



Analiza wariancji i metody klasyfikacyjne

Analysis of variance and classification methods

lecture 2

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Slides: https://indico.ifj.edu.pl/event/271/





Covariance Matrix

• Let X be a p-variate random vector. The covariance matrix of X is defined as:

$$\begin{split} \boldsymbol{\Sigma}_{XX} &= Var(X) = E\{(X-\mu)^T(X-\mu)\} \\ &= \begin{pmatrix} Var(X_1) & Cov(X_1,X_2) & \dots & Cov(X_1,X_p) \\ Cov(X_2,X_1) & Var(X_2) & \dots & Cov(X_2,X_p) \\ \vdots & \vdots & \ddots & \vdots \\ Cov(X_p,X_1) & Cov(X_p,X_2) & \dots & Var(X_p) \end{pmatrix} \end{split}$$

Where:

$$Var(X_i) = E\{(X_i - \mu(X_i)) \cdot (X_i - \mu(X_i))\}$$
$$Cov(X_i, Y_j) = E\{(X_i - \mu(X_i)) \cdot (Y_j - \mu(Y_j))\}$$

Cross-covariance matrix



We define the covariance matrix (cross-covariance) between X and Y to be

$$\begin{split} \boldsymbol{\Sigma}_{XY} &= Cov(\boldsymbol{X}, \boldsymbol{Y}) = E\{(\boldsymbol{X} - \boldsymbol{\mu}_{X})^{T}(\boldsymbol{Y} - \boldsymbol{\mu}_{Y})\} \\ &= \begin{pmatrix} Cov(X_{1}, Y_{1}) & Cov(X_{1}, Y_{2}) & \dots & Cov(X_{1}, Y_{m}) \\ Cov(X_{2}, Y_{1}) & Cov(X_{2}, Y_{2}) & \dots & Cov(X_{2}, X_{m}) \\ \vdots & \vdots & \ddots & \vdots \\ Cov(X_{p}, Y_{1}) & Cov(X_{p}, Y_{2}) & \dots & Cov(X_{p}, Y_{m}) \end{pmatrix} \end{split}$$

Some relations:

$$Cov(AX, BY) = ACov(X, Y)B^T$$
$$Cov(X + a, Y + b) = Cov(X, Y)$$



Trace and variance

Let $A = (a_{ii})$ be a square matrices of dimension $d \times d$.

• The trace of A is the sum of its diagonal elements:

$$tr(A) = \sum_{i} a_{ii}$$



• The mean is the best constant predictor of X in terms of the MSE (shown already at previous lecture): $E(X) = \arg\min E \|X - c\|^2$

 $c \in \mathbb{R}^p$

• The **total variance of X** is defined as the MSE of the mean:

$$\mathbf{E} \mid\mid \mathbf{X} - \mathbf{E}(\mathbf{X}) \mid\mid^{2} = \sum_{i=1}^{N} E(X_{i} - E(X_{i}))^{2} = \sum_{i=1}^{N} Var(X_{i}) = tr(Var(X))$$

 The total variance of X measures the overall variability of the components of X around the mean E(X). Commonly used measure of variability is the standard deviation.





Quadratic Forms

Let A be a symmetric matrix and x a vector.

Definition:

A quadratic form is written as:

$$x^T A x = \sum_i \sum_j a_{ij} x_i x_j$$

Note: it's a quadratic function of x.

- As a function of A, $Var(AX) = A \ Var(X)A^T$ which is a quadratic form in A.
- Quadratic forms are very common in multivariate analysis.
- Example: Chi-squared test is a quadratic form.





Quadratic forms

Forms

$$h_1\left(x_1,x_2,x_3\right) = -x_1^2 + 2x_1x_2 + x_2^2 - 4x_2x_3$$

and

$$h_2\left(x_1, x_2, x_3, x_4\right) = -x_1^2 + 2x_1x_2 + x_2^2 - 4x_2x_3$$

are quadratic forms given by matrices:

$$A_{h_1} = \begin{bmatrix} -1 & 1 & 0 \\ 1 & 1 & -2 \\ 0 & -2 & 0 \end{bmatrix} \quad \text{oraz} \quad A_{h_2} = \begin{bmatrix} -1 & 1 & 0 & 0 \\ 1 & 1 & -2 & 0 \\ 0 & -2 & 0 & 0 \\ 0 & 0 & 0 & 0 \end{bmatrix}$$

These are NOT quadratic forms (think why?):

$$g_1\left(x_1,x_2\right) = x_1^2 + 2x_1x_2 + x_2 \quad \text{oraz} \quad g_2\left(x_1,x_2\right) = x_1^2 + x_2^2 + 1$$

The matrix A is a square matrix and can always be written in a symmetric form. More general: in the complex space it is a Hermitian matrix (or self-adjoint matrix): complex square matrix that is equal to its own conjugate transpose.



Positive/negative semi-definite and positive/negative definite matrix



• A is a positive definite form, if:

$$x^T A x > 0, \quad \forall x \in \mathbb{R}^n \setminus \{0\};$$

• negative definite, if:

$$x^T A x < 0, \quad \forall x \in \mathbb{R}^n \setminus \{0\};$$

• semi-positive definite, if:

$$x^T A x \ge 0, \quad \forall x \in \mathbb{R}^n;$$

• semi-negative definite, if:

$$x^T A x \le 0, \quad \forall x \in \mathbb{R}^n;$$

• If none of the above, the quadratic form is not definite.







Let's take the quadratic form:

$$h_1\left(x_1,x_2,x_3\right) = -x_1^2 + 2x_1x_2 + x_2^2 - 4x_2x_3.$$

Since $h_1(1,0,0) = -1$ and $h_1(0,1,0) = 1$, the quadratic form h_1 is undefinite. Quadratic form

$$h\left(x_{1}, x_{2}\right) = x_{1}^{2} + 2x_{2}^{2}$$

is positive definite, the form:

$$g\left(x_{1}, x_{2}, x_{3}\right) = x_{1}^{2} + 2x_{2}^{2}$$

is semi-positive definite (think why?).

Note that covariance matrices have the following properties: 1)Every covariance matrix is a positive semi-definite matrix. 2)Every positive semi-definite matrix is a covariance matrix.

Methods of testing whether a matrix is for a positive/negative definite

Sylvester criterion:

Quadratic form $h(x) = x^T A x$, where $A = A^T \in \mathbb{R}^{n \times n}$, is:

1) Positive (Negative) Definite when and only when all the leading minors of matrix A are positive (negative):

$$D_j = \left| \begin{array}{ccc} a_{11} & \cdots & a_{1j} \\ \vdots & \ddots & \vdots \\ a_{j1} & \cdots & a_{jj} \end{array} \right| > (<)0, \quad (j = 1, \dots, n) \, ;$$





Sylvester criterion – an example

For a quadratic form

$$h\left(x_{1}, x_{2}, x_{3}\right) = 3x_{1}^{2} + 2x_{1}x_{2} + x_{2}^{2} - 2x_{1}x_{3} + 2x_{3}^{2}$$

we have:

$$h(x_1, x_2, x_3) = \begin{bmatrix} x_1 & x_2 & x_3 \end{bmatrix} \begin{bmatrix} 3 & 1 & -1 \\ 1 & 1 & 0 \\ -1 & 0 & 2 \end{bmatrix} \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$$

Since:

$$D_1 = 3 > 0, \quad D_2 = \left| \begin{array}{ccc} 3 & 1 \\ 1 & 1 \end{array} \right| = 2 > 0, \quad D_3 = \left| \begin{array}{ccc} 3 & 1 & -1 \\ 1 & 1 & 0 \\ -1 & 0 & 2 \end{array} \right| = 3 > 0,$$

therefore a quadratic form h is positive definite.





Determinant

$$\begin{vmatrix} a & b & c & d \\ e & f & g & h \\ i & j & k & l \\ m & n & o & p \end{vmatrix} = a \begin{vmatrix} f & g & h \\ j & k & l \\ n & o & p \end{vmatrix} - b \begin{vmatrix} e & g & h \\ i & k & l \\ m & o & p \end{vmatrix} + c \begin{vmatrix} e & f & h \\ i & j & l \\ m & n & p \end{vmatrix} - d \begin{vmatrix} e & f & g \\ i & j & k \\ m & n & o \end{vmatrix}.$$

The simplest way to express the determinant is by considering the elements in the top row and the respective minors; starting at the left, multiply the element by the minor, then subtract the product of the next element and its minor, and alternate adding and subtracting such products until all elements in the top row have been exhausted.



Eigenvalues and eigenvectors



Let vector $\mathbf{v} > \mathbf{0}$ and let **A** be a d × d matrix.

v is an eigenvector with eigenvalue λ when Av = λv .

- It's typical to normalize the eigenvector to have length 1 (or have it's entries sum to 1).
- Matrix **A** has at most *d* distinct eigenvalues (think about why).
- Eigenvectors with distinct eigenvalues are orthogonal, i.e. $v_1^T v_2 = 0$
- If **A** is a positive definite matrix, then:
 - All of its eigenvalues are real-valued and positive.
 - Its inverse is also positive definite.



Eigendecomposition (spectral decomposition)



- Matrix **A** has n linearly independent eigenvectors $v_1, v_2, ..., v_n$ with associated eigenvalues $\lambda_1, \lambda_2, ..., \lambda_n$.
- Define square matrix Q whose columns are the n linearly independent eigenvectors of A: $Q = \begin{bmatrix} v_1 & v_2 & \cdots & v_n \end{bmatrix}$.
- Since each column of Q is an eigenvector of A: $AQ = \begin{bmatrix} \lambda_1 v_1 & \lambda_2 v_2 & \cdots & \lambda_n v_n \end{bmatrix}$.
- Define matrix Λ : $\Lambda_{ii} = \lambda_i$ and $\Lambda_{ii} = 0$ for $i \neq j$, than $AQ = Q\Lambda$.
- $A = Q \wedge Q^{-1}$ (multiplying both sides by Q^{-1})
- Or $Q^{-1} A Q = \Lambda$
- Matrix A can be decomposed into a matrix composed of its eigenvectors, a diagonal matrix with its eigenvalues along the diagonal, and the inverse of the matrix of eigenvectors.
 - This is called the **eigendecomposition**. Matrix **A** is **diagonalizable**.





Eigendecomposition

• The eigen (spectral) decomposition allows some operations with positive definite matrices to be computed more easily:

 $A^{-1} = P \Lambda^{-1} P^{T}$. $A^{1/2} = P \Lambda^{1/2} P^{T}$.

• Example:

$$\mathbf{A} = \begin{bmatrix} 1 & 0 \\ 1 & 3 \end{bmatrix} \quad \text{can be decomposed into } \begin{bmatrix} -2c & 0 \\ c & d \end{bmatrix}^{-1} \begin{bmatrix} 1 & 0 \\ 1 & 3 \end{bmatrix} \begin{bmatrix} -2c & 0 \\ c & d \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ 0 & 3 \end{bmatrix}, \qquad [c,d] \in \mathbb{R}$$





Eigendecomposition

Scalar $\lambda \in \mathbb{F}$ is called an **eigenvalue** of matrix $A \in \mathbb{F}^{n \times n}$ if exists a non-zero vector $v \in \mathbb{F}^n$, such that:

$$Av = \lambda v;$$

vector v is called an **eigenvector** of an eigenvalue λ . All eigenvalues of matrix A is a spectrum of A and is denoted as $\sigma(A)$.

For matrix $A \in \mathbb{F}^{n \times n}$ the following conditions are equivalent:

- (a) λ is an eigenvalue of A;
- (b) system of equations $(A \lambda I) v = 0$ has a non-zero solution;
- (c) $\det (A \lambda I) = 0.$





Eigendecomposition

For any matrix $A \in F^{n \times n}$ **det(A-\lambdaI)** is a polynomial of degree **n** (characteristic polynomial). The roots of this polynomial are the the eigenvalues.

Example:

Calculate eigenvectors and eigenvalues of matrix:

$$A = \begin{bmatrix} 1 & 2 & 0 \\ 0 & 2 & 0 \\ -2 & -2 & -1 \end{bmatrix}$$

Since $\varphi_A(\lambda) = \det(A - \lambda I) = -(1 - \lambda)(2 - \lambda)(1 + \lambda)$, therefore matrix A has three different eigenvalues: $\lambda_1 = -1, \lambda_2 = 1, \lambda_3 = 2$.

The equation $det(A - \lambda I) = -(1 - \lambda) (2 - \lambda) (1 + \lambda) = 0$ is a characteristic polynomial.



For each eigenvalue we find an eigenvector:



• for $\lambda_1 = -1$ we have:

$$\begin{bmatrix} 2 & 2 & 0 \\ 0 & 3 & 0 \\ -2 & -2 & 0 \end{bmatrix} \begin{bmatrix} x \\ y \\ z \end{bmatrix} = \begin{bmatrix} 2x+2y \\ 3y \\ -2x-2y \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix}$$

so we get $(x,y,z)=(0,0,t)\,,\ t\in\mathbb{R};$ an example eigenvector $v_{\lambda_1}=(0,0,1)^T\,;$

• for $\lambda_2 = 1$ we get:

$$\begin{bmatrix} 0 & 2 & 0 \\ 0 & 1 & 0 \\ -2 & -2 & -2 \end{bmatrix} \begin{bmatrix} x \\ y \\ z \end{bmatrix} = \begin{bmatrix} 2y \\ y \\ -2x - 2y - 2z \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix}$$

so we get $(x,y,z)=(t,0,-t)\,,\ t\in\mathbb{R};$ example eigenvector is $v_{\lambda_2}=(1,0,-1)^T\,;$

• for $\lambda_3 = 2$ we get:

$$\begin{bmatrix} -1 & 2 & 0 \\ 0 & 0 & 0 \\ -2 & -2 & -3 \end{bmatrix} \begin{bmatrix} x \\ y \\ z \end{bmatrix} = \begin{bmatrix} -x+2y \\ 0 \\ -2x-2y-3z \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix}$$

we get $\left(x,y,z\right)=\left(2t,t,-2t\right),\,t\in\mathbb{R};$ eigenvector $v_{\lambda_{3}}=\left(2,1,-2\right)^{T}.$





Multivariate Normal Distribution

 MVN is generalization of univariate normal distribution (gausian).





Johann Karl Friedrich Gauss (1777 – 1855)





Multivariate Normal Distribution

• We assume that the population mean is $\mu = E(X)$ and $\Sigma = Var(X) = E[(X - \mu)(X - \mu)^T]$, then:

$$\mu = \begin{pmatrix} \mu_1 \\ \mu_2 \\ \vdots \\ \mu_p \end{pmatrix}$$

and

$$\Sigma = \begin{pmatrix} \sigma_1^2 & \sigma_{12} & \dots & \sigma_{1p} \\ \sigma_{21} & \sigma_2^2 & \dots & \sigma_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{p1} & \sigma_{p2} & \dots & \sigma_p^2 \end{pmatrix}$$





Central limit theorem (CLT)

• One of the **most important theorems** of the statistics – this is why we observe in nature mostly Gaussian distributions.

Let $\{X_1, ..., X_n\}$ be a random sample of size n of independent and identically distributed random variables drawn from a distribution of expected value given by μ and finite variance given by σ^2 . CLT states that as n gets larger, the distribution of the difference between the sample average and its limit μ , when multiplied by the factor \sqrt{n} , approximates the normal distribution with mean 0 and variance σ^2 . For large enough n, the distribution of average is close to the normal distribution with mean μ and variance σ^2/n .







Central limit theorem (CLT)



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Central limit theorem (CLT)

For the first time CLT for binomial distributions was postulated in the second edition of the *"The Doctrine of Chances"* by Abraham de Moivre'a, published in 1738. It was forgotten for over 80 years, and in 1812 Pierre-Simon Laplace proved CLT for the binomial distributions.

CLT in the version of Lindeberg & Levy was published in 1920'ties, however independently it was proven earlier by Aleksandr Lyapunov in 1901.



DOCTRINE

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Abraham de Moivre (1667 – 1754)





Proof of classical CLT

- Characteristic function of a random variable defines its probability distribution (if a random variable admits a probability density function, then the characteristic function is the Fourier transform of the probability density function).
- Assume $\{X_1, ..., X_n\}$ are independent and identically distributed random variables, each with mean μ and finite variance σ^2 . The sum $X_1 + ... + X_n$ has mean $n\mu$ and variance $n\sigma^2$. The random variable:

$$Z_n = \frac{X_1 + \dots + X_n - n\mu}{\sqrt{n\sigma^2}} = \sum_{i=1}^n \frac{X_i - \mu}{\sqrt{n\sigma^2}} = \sum_{i=1}^n \frac{1}{\sqrt{n}} Y_i, \qquad \qquad Y_i = \frac{X_i - \mu}{\sigma}$$

• The characteristic function of Zn is given by:

$$arphi_{Z_n}(t) \ = \ arphi_{\sum_{i=1}^n rac{1}{\sqrt{n}}Y_i}(t) \ = \ arphi_{Y_1}\!\left(rac{t}{\sqrt{n}}
ight)arphi_{Y_2}\!\left(rac{t}{\sqrt{n}}
ight)\cdots arphi_{Y_n}\!\left(rac{t}{\sqrt{n}}
ight) \ = \ \left[arphi_{Y_1}\!\left(rac{t}{\sqrt{n}}
ight)
ight]^n,$$

• since all of the Y_i are identically distributed (zero mean, σ =1).

$$\varphi_{Y_1}\!\left(\frac{t}{\sqrt{n}}\right) \ = \ 1 - \frac{t^2}{2n} + o\!\left(\frac{t^2}{n}\right), \quad \left(\frac{t}{\sqrt{n}}\right) \to 0 \qquad \text{ Taylor's theorem}$$

"An Introduction to Stochastic Processes in Physics", Don S. Lemons

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Proof of classical CLT

- Since $e^{x} = \lim(1 + x/n)^{n}$, the characteristic function of Z_{n} equals: $\varphi_{Z_{n}}(t) = \left(1 - \frac{t^{2}}{2n} + o\left(\frac{t^{2}}{n}\right)\right)^{n} \to e^{-\frac{1}{2}t^{2}}, \quad n \to \infty.$
- All of the higher order terms vanish in the limit $n \rightarrow \infty$.
- The right hand side equals the characteristic function of a standard normal distribution N(0,1) \rightarrow in the limit of n $\rightarrow \infty$ the Z_n \rightarrow N(0,1)







Rank 100 Tiger



Rank 30 Tiger



Rank 10 Tiger



Rank 200 Tiger



Rank 50 Tiger



Rank 20 Tiger



Rank 3 Tiger









Singular value decomposition is a method of decomposing a matrix into three other matrices:

SVD

$$A = USV^T$$

Where:

$$AA^TU = US^2$$

A is an m × n ma U is an m × n orthogonal matrix S is an n × n diagonal matrix V is an n × n orthogonal matrix

Orthogonal matrices:

 $U^T U = V V^T = I$

https://blog.statsbot.co/singular-value-decomposition-tutorial-52c695315254







$$A = USV^T \qquad \qquad a_{ij} = \sum_{k=1}^n u_{ik} s_k v_{jk}$$

The variables, $\{s_i\}$, are called singular values and are normally arranged from largest to smallest:

$$s_{i+1} \le s_i$$

The columns of U are called left singular vectors, while those of V are called right singular vectors.





Using orthogonality property we get:

$$A = USV^T$$

 $AA^{T}U = US^{2}$ $A^{T}AV = VS^{2}$

The standard procedure (or eigenvalue calculator) can be used to solve these equations and find the U, V and S^2 .

SVD

Applications of the SVD

the SVD of a 32-times-32 digital image A is computed

the activities are lead by Prof. Per Christian Hansen.





Applications of the SVD



A



14.10.2019

 A_{5}

<u>G.</u>Strang " at first you see nothing, and wolderly you recognize everything."







- 1) Write a code (python, R) which performs eigendecomposition of a given symmetric matrix.
 - Code it yourself
 - Try maybe numpy.roots(p) to find the roots
 - Test against numpy.linalg.eig()
 - If you want to play more try to visualize the linear transformation (see "BONUS: visualizing linear transformations" in

https://hadrienj.github.io/posts/Deep-Learning-Book-Series-2.7-Eigendecomposition/







Exercises

• Apply Singular Value Decomposition to a photograph:

numpy.linalg.svd











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Exercises



• Write a script showing, that the CLT works :)

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