

# Statistical issues for Higgs Physics

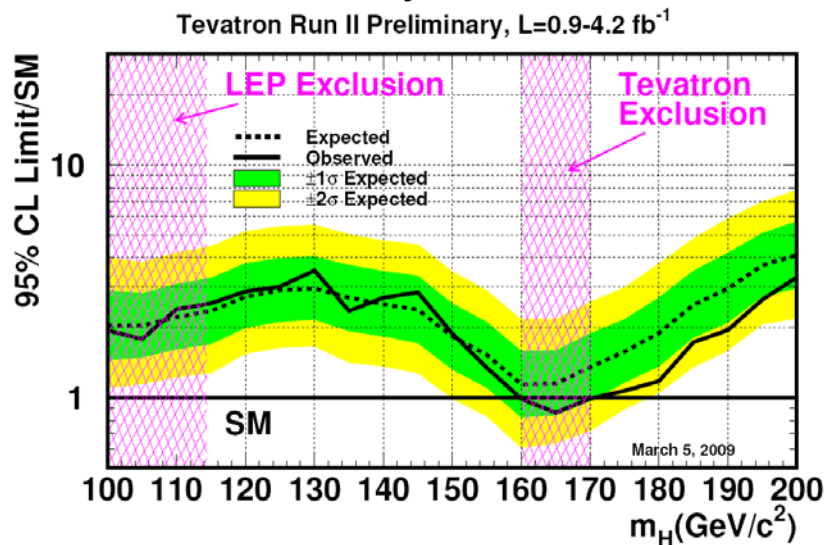
Eilam Gross, Weizmann Institute of Science

Acknowledgements: Louis Lyons & Ofer Vitells



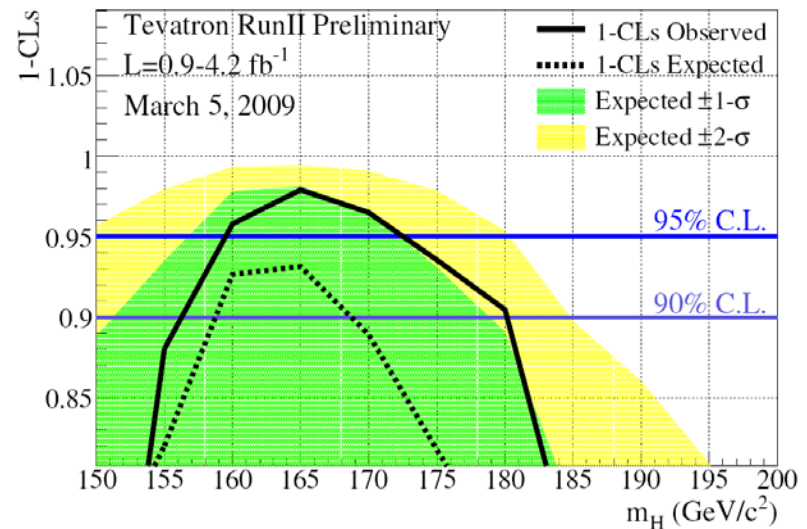
# Combined CDF and DØ Upper Limits on Standard Model Higgs-Boson Production with up to 4.2 fb<sup>-1</sup> of Data

## Bayesian



FERMILAB-PUB-09-060-E  
CDF Note 9713  
DØ Note 5889

## Modified Frequentist - CLs

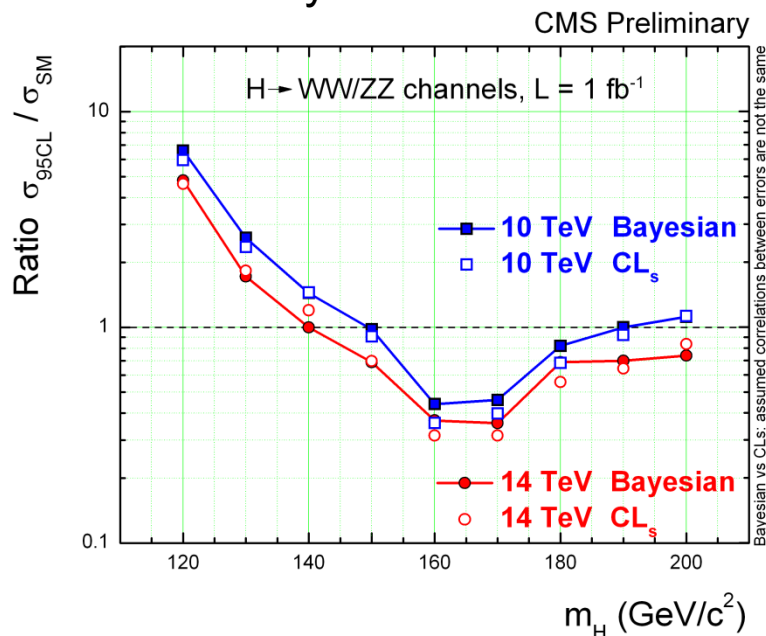


FERMILAB-PUB-09-060-E  
CDF Note 9713  
DØ Note 5889



# CMS & ATLAS Higgs Prospects

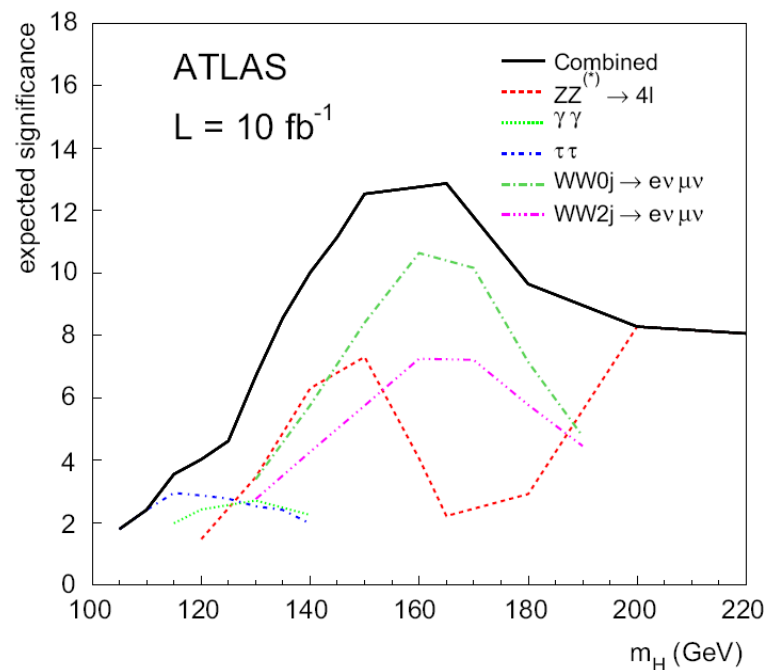
## Bayesian



CMS

Gregory Schott SUSY 2009

## Profile Likelihood



ATLAS

CERN-OPEN-2008-020

December 2008



# Discovery vs Exclusion

- Higgs statistics is about testing one hypothesis against another hypothesis
  - One hypothesis is the Standard Model with no Higgs Boson ( $H_0$ )
  - Another hypothesis is the SM with a Higgs boson with a specific mass  $m_H$  ( $H_1$ )
- Rejecting the No-Higgs ( $H_0$ ) hypothesis  $\rightarrow$  DISCOVERY
- Rejecting the Higgs hypothesis ( $H_1$ )  $\rightarrow$  EXCLUDING the Higgs



# DISCOVERY



# The Discovery Case Study

- We assume a Gaussian “Higgs” signal (**s**) on top of a Rayleigh shaped background (**b**)

NOTE:  $b=b(\theta)$

$$m = \tau b$$

$$n = \mu s(m_H) + b$$

- The signal strength is  $\mu$

$$\langle n \rangle = \mu s + b$$

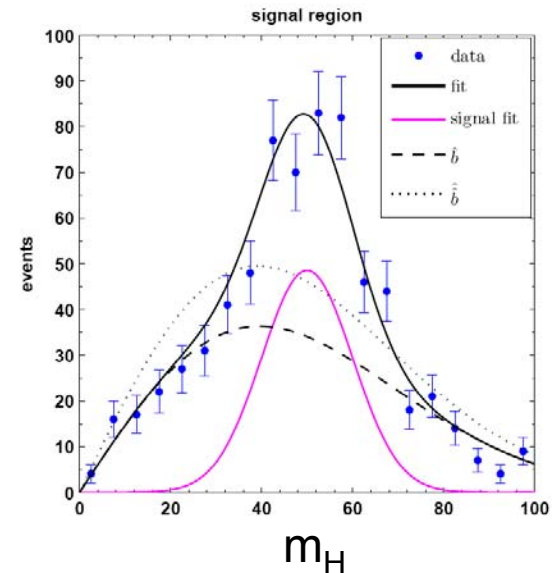
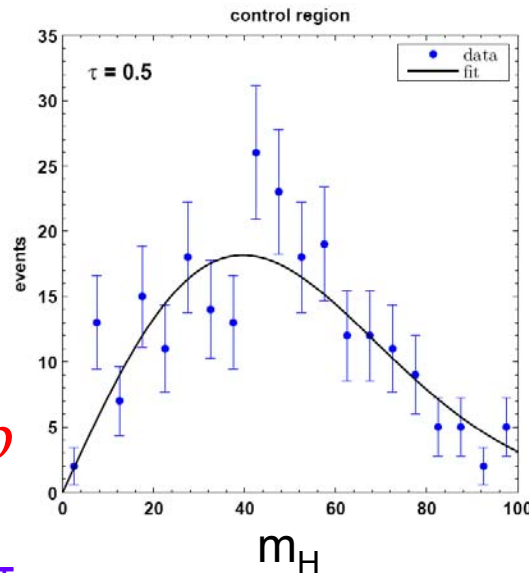
- $\mu=1$ , SM Higgs
- $\mu=0$ , SM without Higgs

- Two hypothetical measurements

- Data  $n \sim \mu s(m_H) + b$

- BG control sample scaled to the expected BG via a factor  $\tau$

$$m \sim \tau b$$

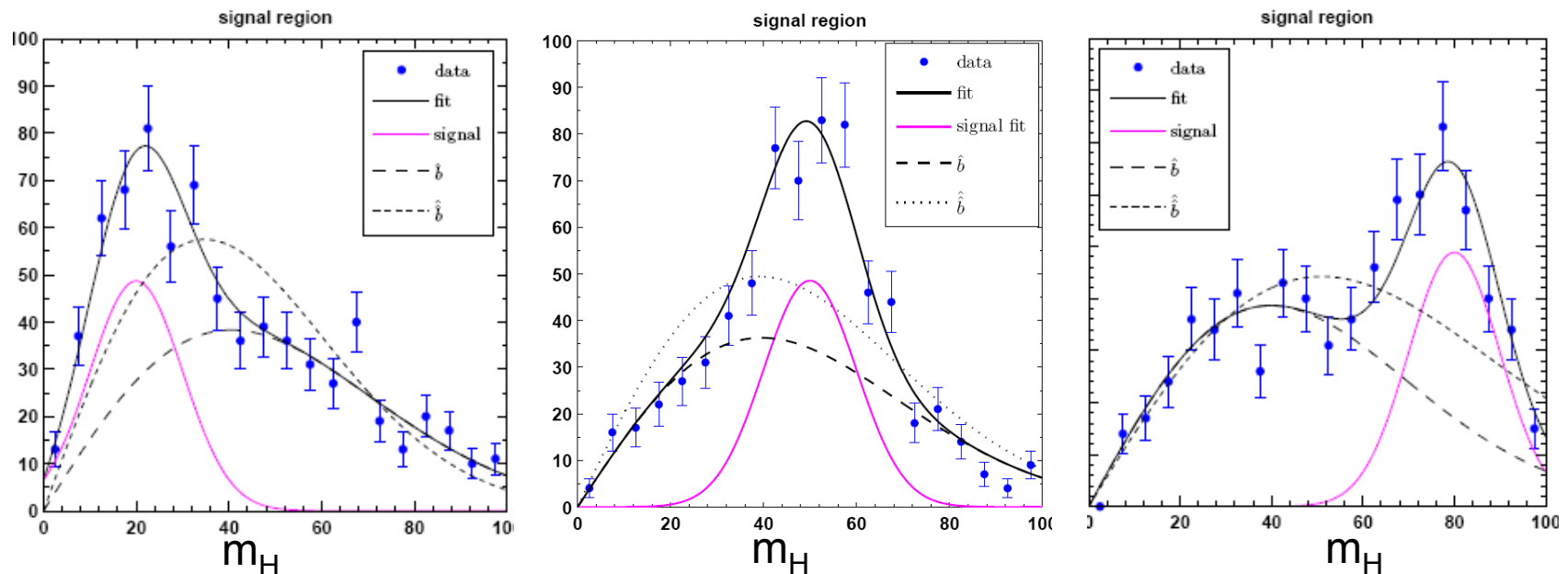


$$L(n, m | s + b(\theta)) = \prod_{i=1}^{nbins} \text{Poisson}(n_i; s_i + b_i) \cdot \text{Poisson}(m_i; \tau b_i)$$



# The Discovery Case Study

Note, in this example, the signal towards the end of the background mass distribution ( $m_H=20,80$ ) is better separated from the signal near the middle ( $m_H=50$ ).



# CL motivation-

## The Neyman-Pearson Lemma

- When performing a hypothesis test between two simple hypotheses,  $H_0$  and  $H_1$ , **the Likelihood Ratio test**, which rejects  $H_0$  in favor of  $H_1$ ,  
**is the most powerful test**

of size  $\alpha$  for a threshold  $\eta$   $\Lambda(x) = \frac{L(H_1 | x)}{L(H_0 | x)} \leq \eta, P(\Lambda(x) \leq \eta | H_0) = \alpha$

- Define a **test statistic**  $\Lambda = \frac{L(H_1)}{L(H_0)}$
- **Note:** Likelihoods are functions of the data,  $\Lambda(x) = \frac{L(H_1 | x)}{L(H_0 | x)}$  even though we often not specify it explicitly





# The frequentist LR (CL) method

- Define a test statistics

$$\Lambda = \frac{L(H_1)}{L(H_0)} = \frac{L(n, m | s + b(\theta))}{L(n, m | b(\theta))}$$

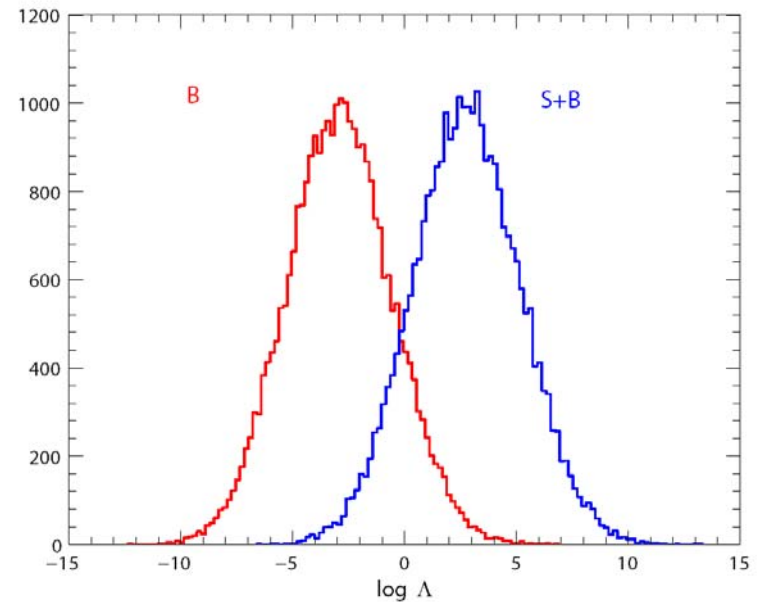


# The frequentist LR (CL) method

- Define a test statistics

$$\Lambda = \frac{L(H_1)}{L(H_0)} = \frac{L(n, m | s + b(\theta))}{L(n, m | b(\theta))}$$

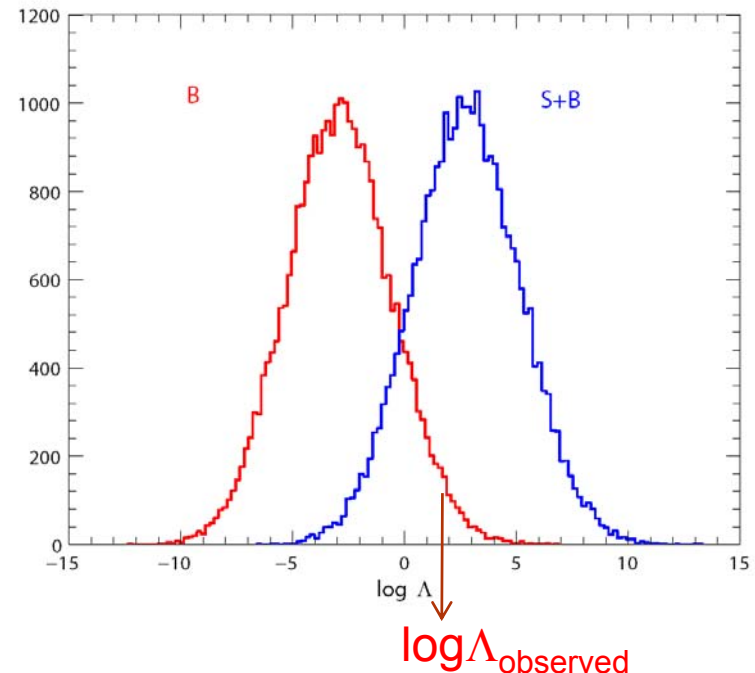
- Use MC to generate the pdf of  $\Lambda$  under  $H_0$ (B only) and  $H_1$  (S+B)



# The frequentist LR (CL) method

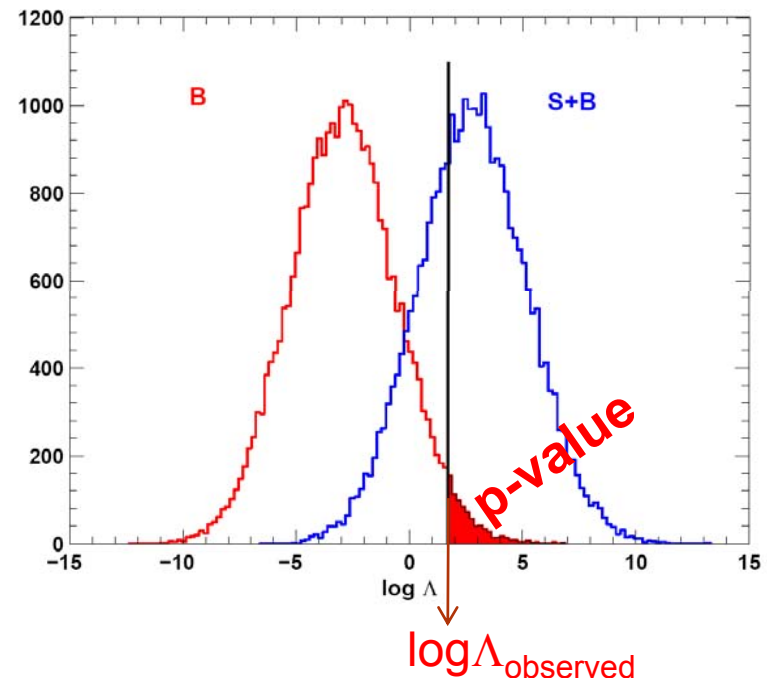
- Define a test statistics
- Use MC to generate the pdf of  $\Lambda$  under  $H_0$  (B only) and  $H_1$  (S+B)
- Let  $\Lambda_{\text{obs}}$  be a result of **one experiment** (LHC)

$$\Lambda = \frac{L(H_1)}{L(H_0)} = \frac{L(n, m | s + b(\theta))}{L(n, m | b(\theta))}$$



# The frequentist LR (CL) method

- Define a test statistics
$$\Lambda = \frac{L(H_1)}{L(H_0)} = \frac{L(n, m | s + b(\theta))}{L(n, m | b(\theta))}$$
- Use MC to generate the pdf of  $\Lambda$  under  $H_0$ (B only) and  $H_1$  (S+B)
- Let  $\Lambda$  be a result of **one experiment** (LHC)
- The p-value is the probability to get an observation which is less B-like than the observed one
- If the result of the experiment (LHC) yields a  $p\text{-value} < 2.8 \cdot 10^{-7}$  a  $5\sigma$  discovery is claimed
- NOTE: the p-value can be interpreted as a frequency  $\rightarrow$  this is a **frequentist** approach

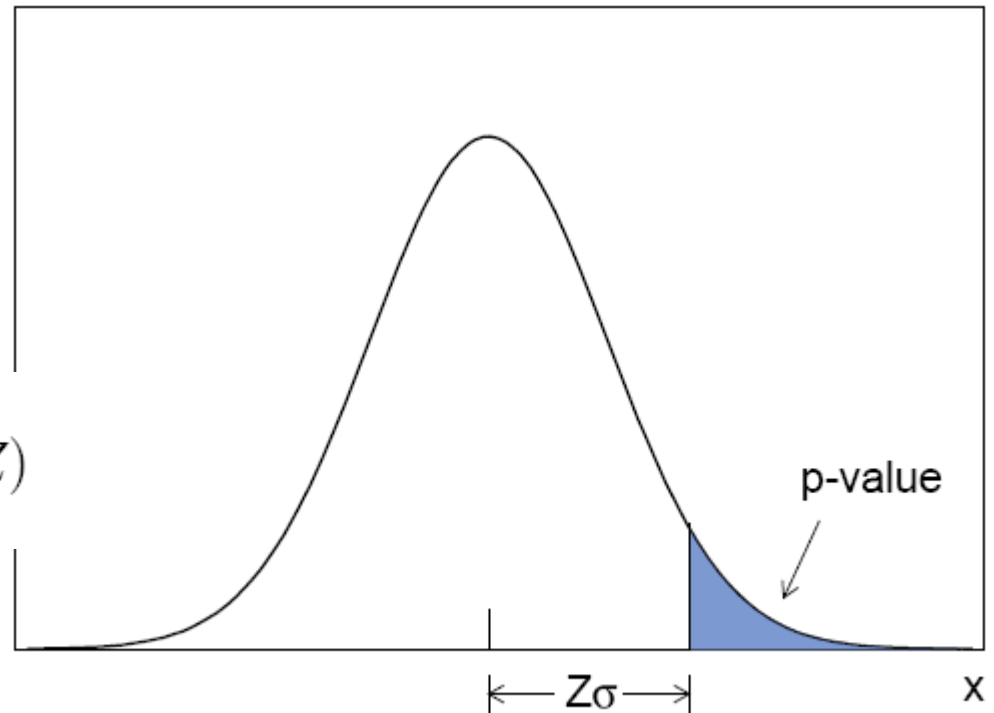


# From p-values to Gaussian significance

It is a custom to express the p-value as the significance associated to it, had the pdf were Gaussians

$$p = \int_Z^{\infty} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx = 1 - \Phi(Z)$$

$$Z = \Phi^{-1}(1 - p)$$



A significance of  $Z = 5$  corresponds to  $p = 2.87 \times 10^{-7}$

# Systematics

- Normally, the background,  $b(\theta)$ , has an uncertainty which has to be taken into account. In this case  $\theta$  is called a nuisance parameter (which we associate with background systematics)
- How can we take into account the nuisance parameters?
- One way: marginalize them (integrate them out using priors)  
 $\rightarrow$  the Hybrid CL (mix frequentist and Bayesian approach)

R.D. Cousins and V.L. Highland. Incorporating systematic uncertainties into an upper limit. *Nucl. Instrum. Meth.*, A320:331–335, 1992.

$$\Lambda_{Hybrid} = \frac{\int L(n, m | s + b(\theta)) \pi(\theta) d\theta}{\int L(n, m | b(\theta)) \pi(\theta) d\theta}$$

prior  $\nearrow$

- Another way is profiling via the MLEs:

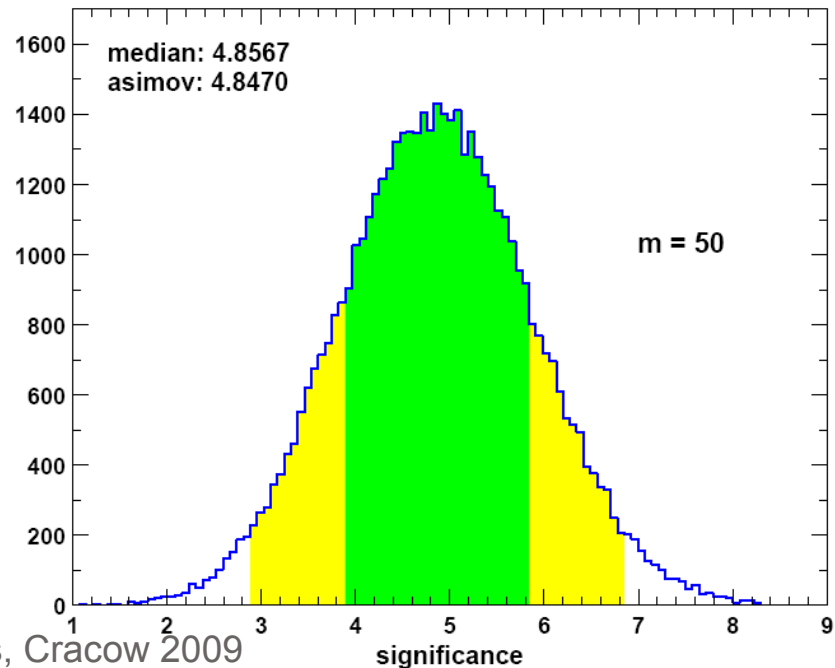
$$\Lambda_{PL} = \frac{L(H_1)}{L(H_0)} = \frac{L\left(n, m | s + b(\hat{\theta}_{s+b})\right)}{L\left(n, m | b(\hat{\theta}_b)\right)}$$



# The Profiled CL way

$$\Lambda_{PL} = \frac{L(H_1)}{L(H_0)} = \frac{L\left(n, m \mid s + b(\hat{\theta}_{s+b})\right)}{L\left(n, m \mid b(\hat{\theta}_b)\right)}$$

- The median significance can be obtained with the one Asimov data set  
 $n \sim s+b, m \sim b$



ATLAS, CERN – Open 2008-029  
Cowan, Cranmer, E.G., Vitells, in preparation

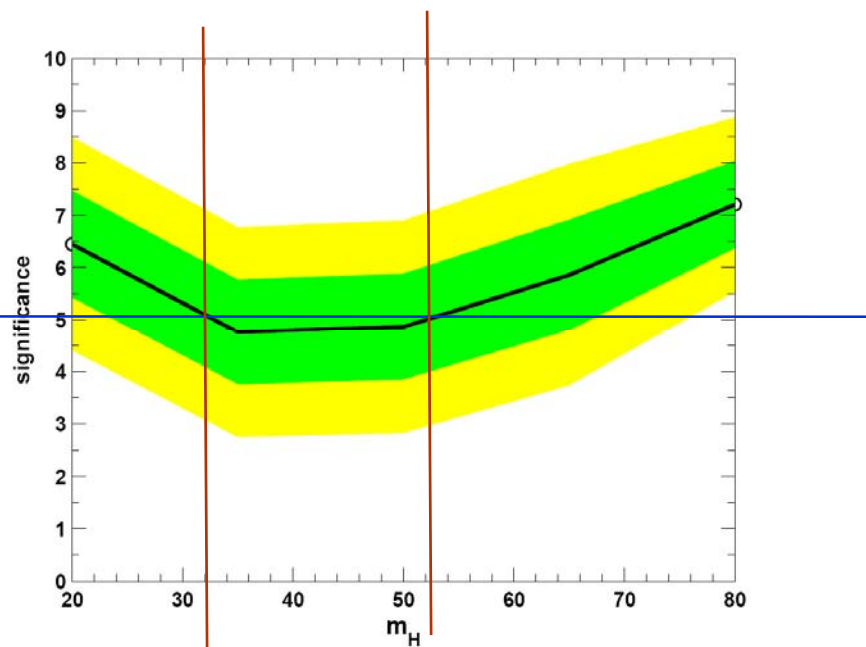


# The Profiled CL way

$$\Lambda_{PL} = \frac{L(H_1)}{L(H_0)} = \frac{L\left(n, m \mid s + b(\hat{\theta}_{s+b})\right)}{L\left(n, m \mid b(\hat{\theta}_b)\right)}$$

In this example a Higgs with a mass  $m_H < 32$  or  $m_H > 52$  is expected to be discovered, i.e.

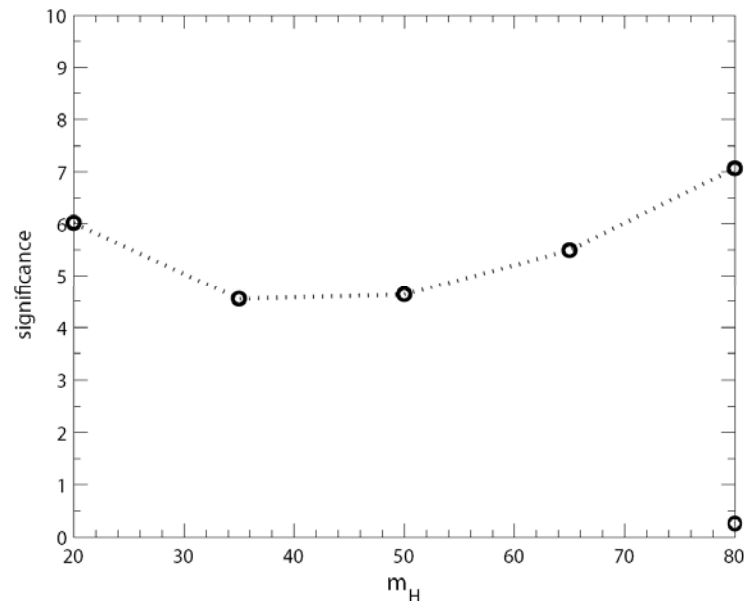
if the Higgs exists in this mass range it will be discovered  $>50\%$  of hypothetical LHC experiments





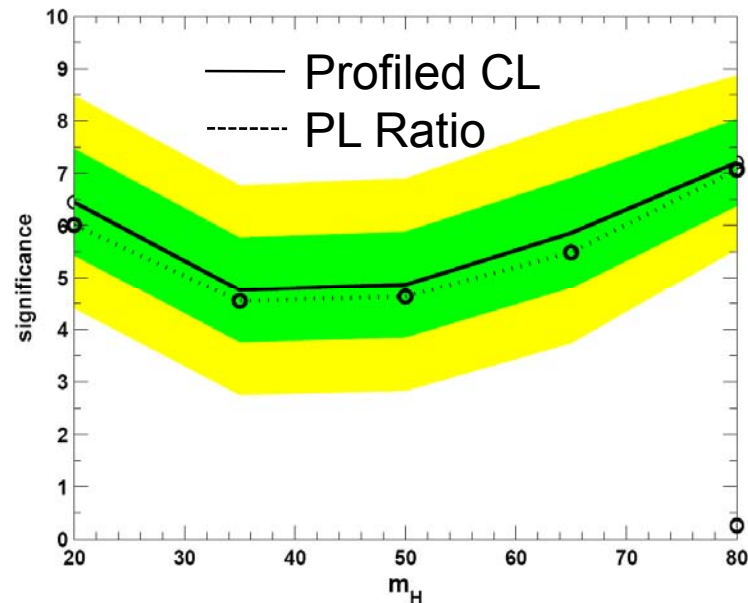
# The Profile Likelihood Ratio

- PL Ratio:  $\lambda(\mu) = \frac{L\left(\mu \cdot s + b(\hat{\theta}_{(\mu=1)})\right)}{L\left(\hat{\mu} \cdot s + b(\hat{\theta})\right)}$
- Test the null  $H_0$  hypothesis  $\lambda(0) = \frac{L\left(b(\hat{\theta}_{(\mu=1)})\right)}{L\left(\hat{\mu} \cdot s + b(\hat{\theta})\right)}$



# The frequentist Profile Likelihood Ratio vs Profiled CL

- Why using a method with a slightly lower sensitivity?



# The Profile Likelihood Ratio & Wilks theorem

- PL Ratio:  $\lambda(\mu) = \frac{L\left(\mu \cdot s + b(\hat{\theta}_{(\mu=1)})\right)}{L\left(\hat{\mu} \cdot s + b(\hat{\theta})\right)}$

- Test the null  $H_0$  hypothesis

$$\lambda(0) = \frac{L\left(b(\hat{\theta}_{(\mu=1)})\right)}{L\left(\hat{\mu} \cdot s + b(\hat{\theta})\right)}$$

- Test statistics  $q_0 = -2 \ln \lambda(0)$

- For B-only experiments

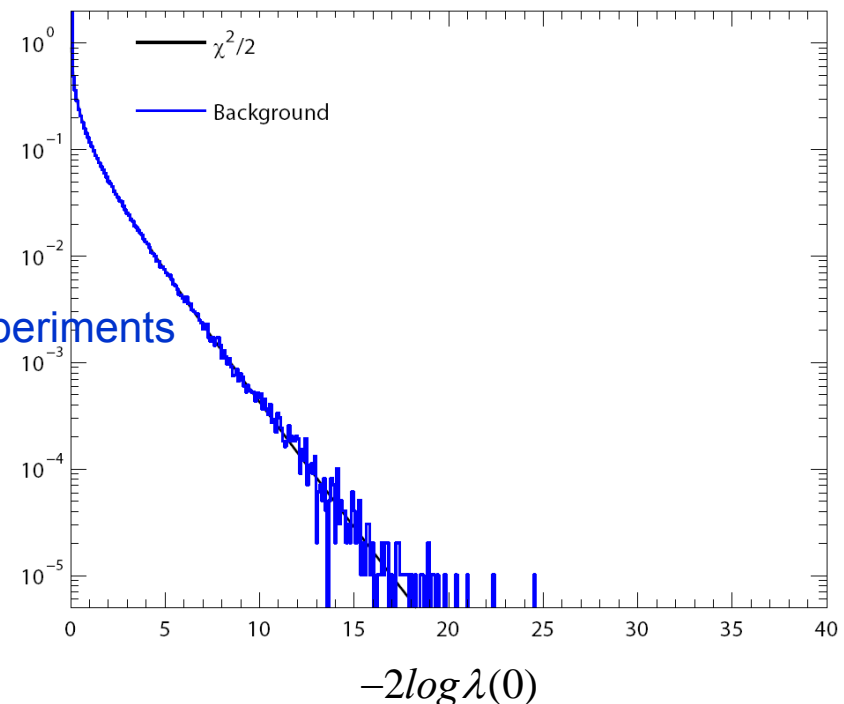
$$\hat{\mu} \sim 0, q_0 \sim -2 \ln 1 \sim 0$$

Wilks theorem:

$q_0 = -2 \log \lambda(0)$  distributes as  $\chi^2$  for b-only experiments

$$\left( \hat{\mu} > 0 \rightarrow \frac{1}{2} \chi^2 \right)$$

Using Wilks theorem there is no need to generate a Billion of events to get the BG pdf and test a  $5\sigma$  sensitivity



# The Profile Likelihood Ratio

- PL Ratio:  $\lambda(\mu) = \frac{L(\mu \cdot s + b(\hat{\theta}_{(\mu=1)}))}{L(\hat{\mu} \cdot s + b(\hat{\theta}))}$

- Test the null  $H_0$  hypothesis  $\lambda(0) = \frac{L(b(\hat{\theta}_{(\mu=1)}))}{L(\hat{\mu} \cdot s + b(\hat{\theta}))}$

- Test statistics  $q_0 = -2 \ln \lambda(0)$

- For B-only experiments

$\hat{\mu} \sim 0$ ,  $q_0 = -2 \ln 1 \sim 0$   
and distributes as  $\chi^2$

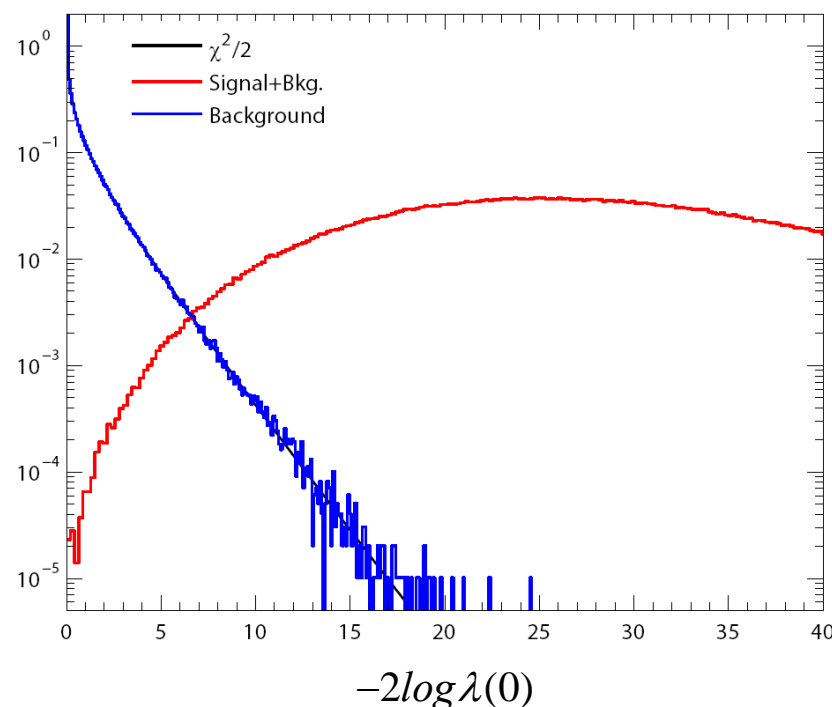
- For S+B experiments

$\hat{\mu} \sim 1$ ,  $q_0 = -2 \ln 0 \gg 1$

- The significance is given to a very good approximation by

$$Z(x) = \sqrt{q_0(x)} = \sqrt{-2 \ln \lambda(0 | x)}$$

and is dependent on the data x



# The Profile Likelihood Ratio

- PL Ratio:  $\lambda(\mu) = \frac{L(\mu \cdot s + b(\hat{\theta}_{(\mu=1)}))}{L(\hat{\mu} \cdot s + b(\hat{\theta}))}$
- Test the null  $H_0$  hypothesis  $\lambda(0) = \frac{L(b(\hat{\theta}_{(\mu=1)}))}{L(\hat{\mu} \cdot s + b(\hat{\theta}))}$

• Test statistics  $q_0 = -2 \ln \lambda(0)$

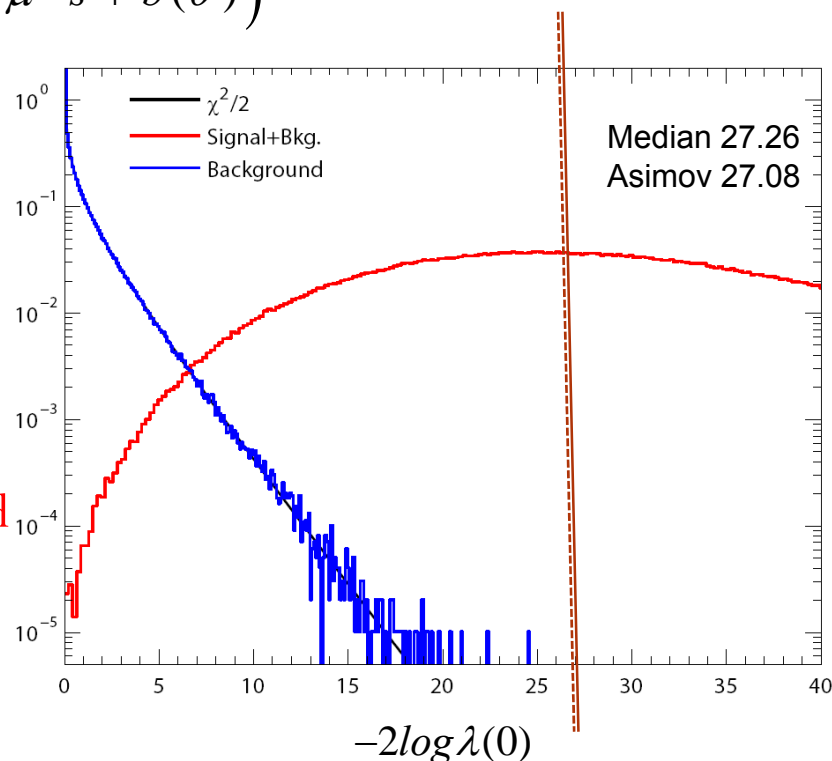
• For B-only experiments

$\hat{\mu} \sim 0$ ,  $q_0 = -2 \ln 1 \sim 0$   
and distributes as  $\chi^2$

• For S+B experiments

$\hat{\mu} \sim 1$ ,  $q_0 = -2 \ln 0 \gg 1$

• The median significance can be obtained  
by the median of the  $q_0 = -2 \ln \lambda(0)$   
distribution or  
using the Asimov data ( $x \sim S+B$ )



# The Bayes way

- Use the observed single data set  $x$  with priors to derive the posterior probability  $P(H_1 | x)$  based on Bayes' theorem

$$P(H_i | x) = \frac{L(x | H_i) \pi(H_i)}{P(x)}$$

- To claim a strong evidence of  $H_1$  over  $H_0$  (a discovery) define the Bayes factor  $B_{10}$  as the ratio of the posterior to prior odds

$$B_{10} = \frac{P(H_1 | x) / P(H_0 | x)}{P(H_1) / P(H_0)}$$

$$B_{10}(\mu) = \frac{\int L(n, m | \mu \cdot s + b(\theta)) \pi(\theta) d\theta}{\int L(n, m | (\mu = 0) \cdot s + b(\theta)) \pi(\theta) d\theta}$$

$$B_{10} = \frac{\iint L(n, m | \mu \cdot s + b(\theta)) \pi(\mu) \pi(\theta) d\theta d\mu}{\int L(n, m | (\mu = 0) \cdot s + b(\theta)) \pi(\theta) d\theta}$$



# PL vs Bayesian

- Using the saddle point approximation one can approximate (see appendix)

$$\log B_{10} \approx \frac{-2 \log \lambda(0)}{2} + C; \quad \log B_{10} \approx \frac{Z^2}{2} + C$$

and get a **rough estimation** for  $\mathbf{B}_{10}$  (neglecting  $C$ )

$Z$	$\sim \mathbf{B}_{10}$	
1	1.6	No evidence
2	7.3	Weak evidence
3	90	Evidence
5	26800	Discovery



# Look Elsewhere Effect

- To establish a discovery we try to reject the background only hypothesis  $H_0$  against the alternate hypothesis  $H_1$
- $H_1$  could be
  - A Higgs Boson with a specified mass  $m_H$
  - A Higgs Boson at some mass  $m_H$  in the search mass range
- The look elsewhere effect deals with the floating mass case

Let the Higgs mass,  $m_H$ , and  
the signal strength  $\mu$   
be 2 parameters of interest

$$\lambda(\mu, m_H) = \frac{L(\mu, m_H, \hat{\hat{b}})}{L(\hat{\mu}, \hat{m}_H, \hat{\hat{b}})}$$

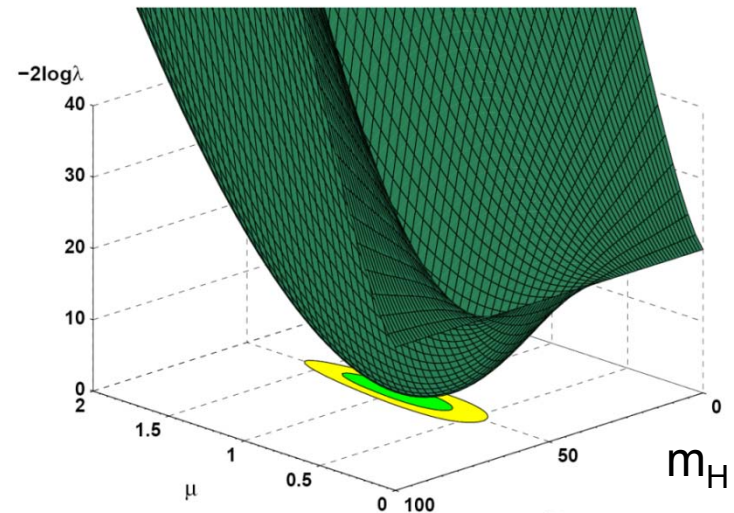
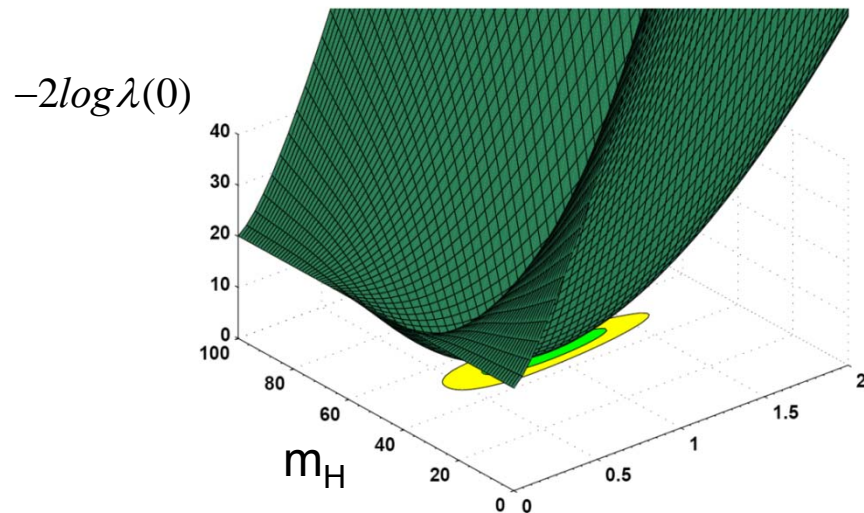




# Look Elsewhere Effect

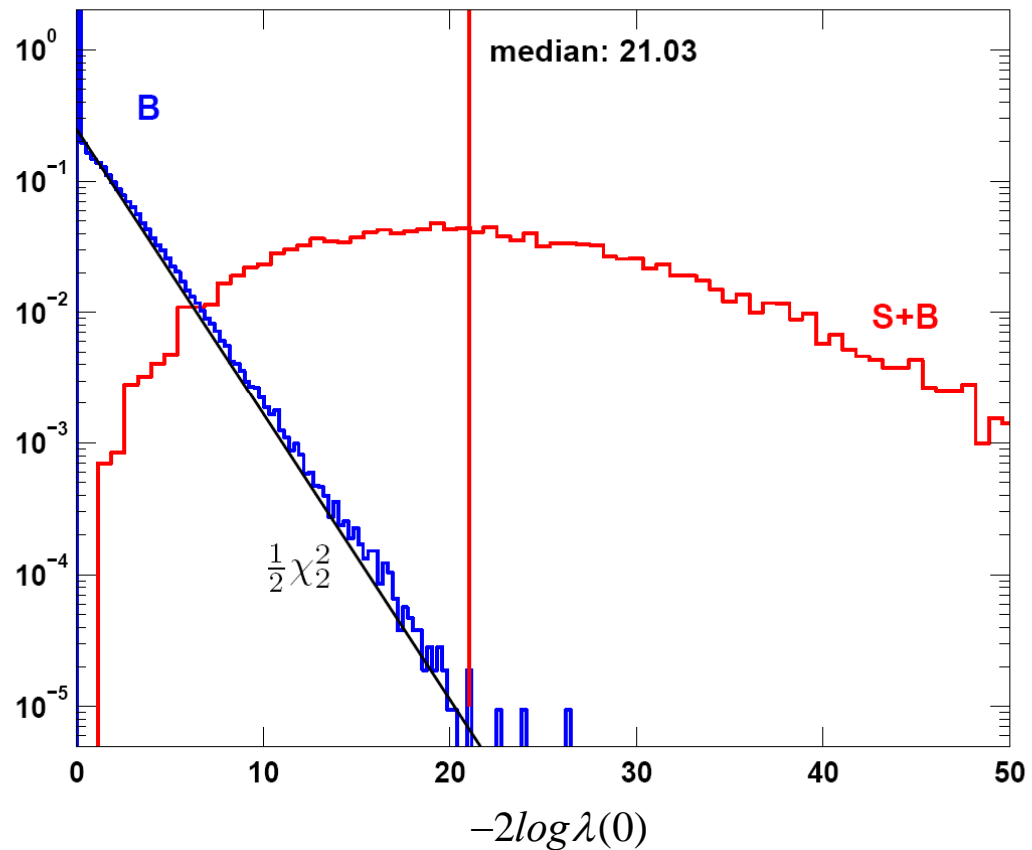
2 parameters of interest: the signal strength  $\mu$  and the Higgs mass  $m_H$

$$\lambda(\mu, m_H \mid n = s_{(m_H=50)} + b, m = \tau b) = \frac{L(\mu, m_H, \hat{b})}{L(\hat{\mu}, \hat{m}_H, \hat{b})}$$



# Look Elsewhere Effect

- Letting the Higgs mass float Wilks' theorem tells us that the background-only experiments will distribute as a  $\chi^2_2$
- The median sensitivity is given by the corresponding p-value

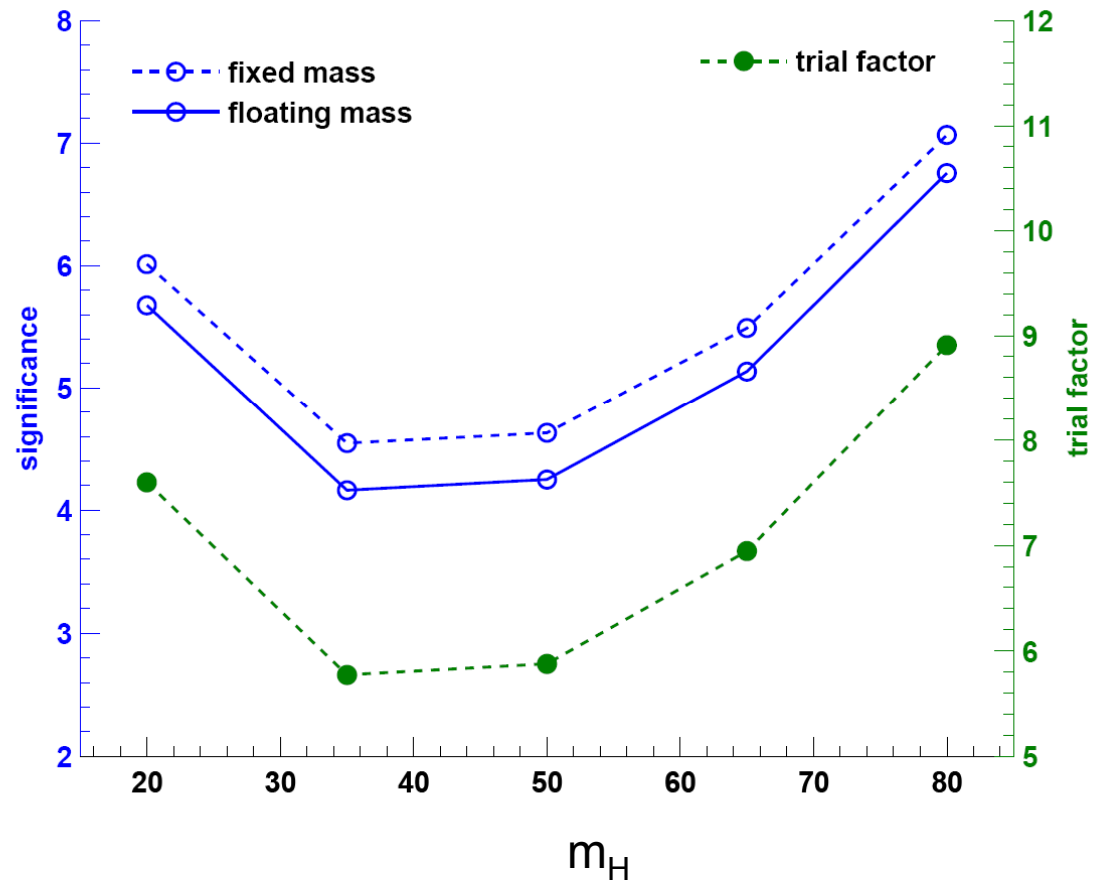


# Look Elsewhere Effect

$$\text{trial factor} = \frac{p_{\text{float}}}{p_{\text{fix}}}$$

Back of the envelope:

$$\text{trial factor} = \frac{\text{range}}{\text{resolution}} = \frac{\Gamma_m}{\sigma_m}$$



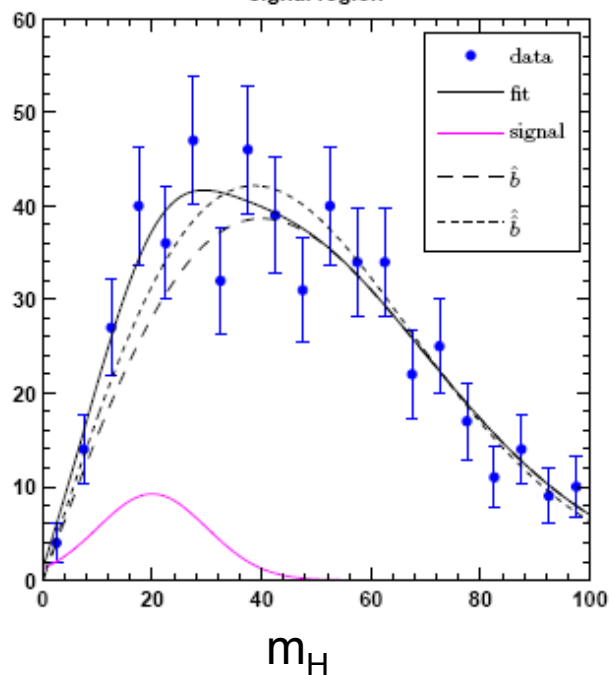
# EXCLUSION



# Exclusion Case Study

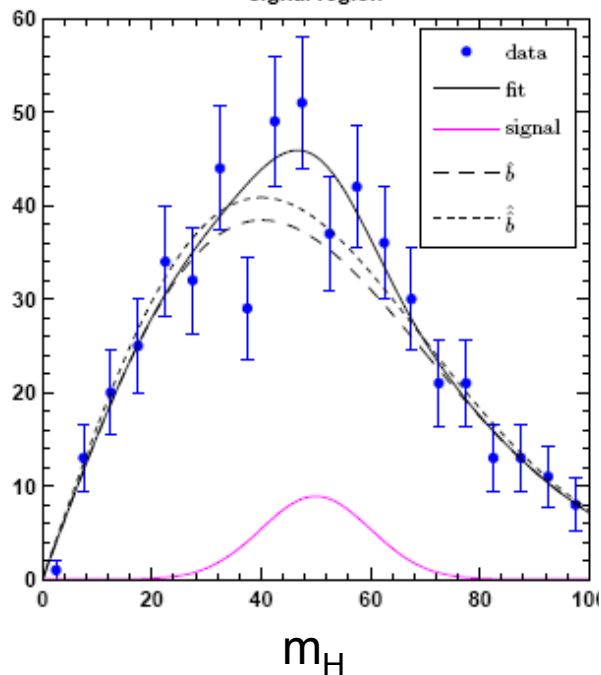
M=20

signal region



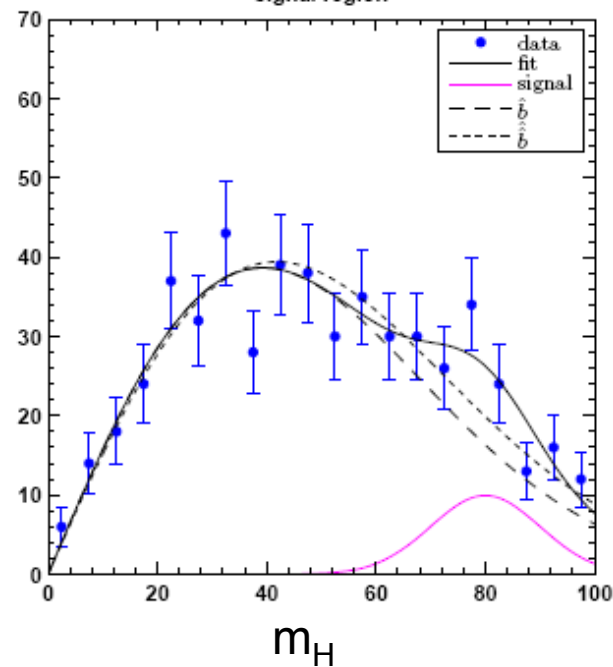
M=50

signal region



M=80

signal region



# Profile Likelihood Ratio

Test the  $S(m_H)+b$  hypothesis  
i.e. test the  $\mu=1$  hypothesis

$$\lambda(\mu = 1) = \frac{L\left(s + b(\hat{\theta}_{(\mu=1)})\right)}{L\left(\hat{\mu} \cdot s + b(\hat{\theta})\right)}$$

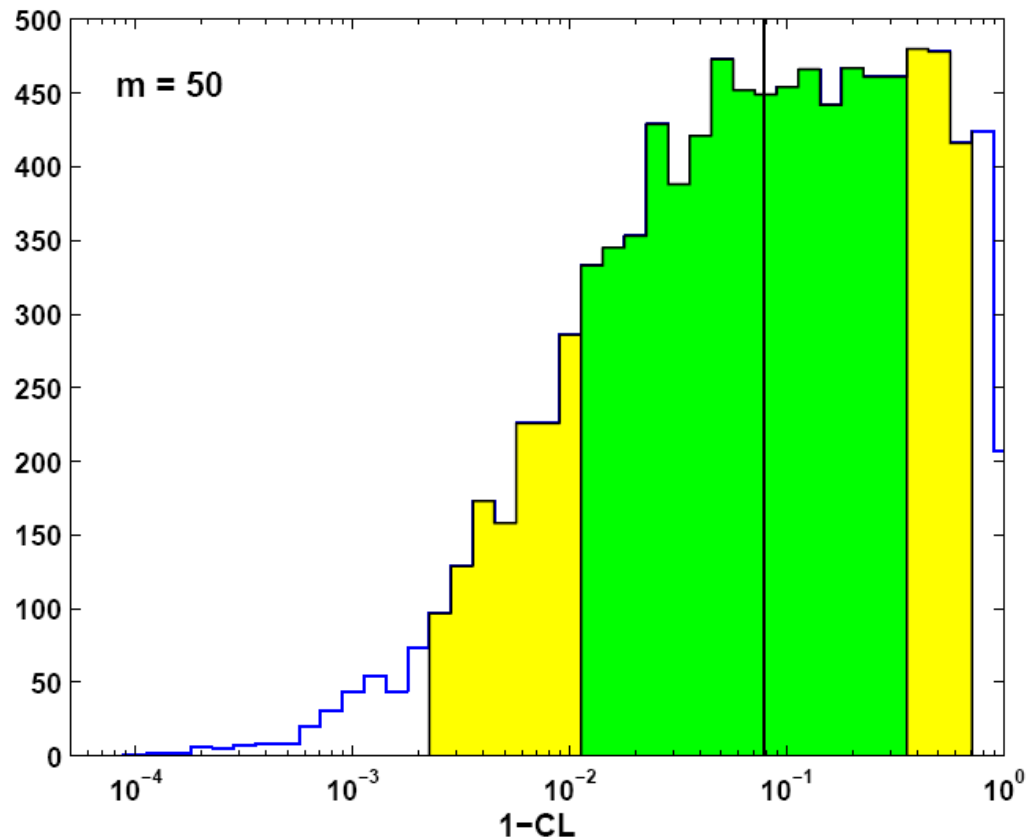
$$q_1 = -2 \ln \lambda(\mu = 1)$$

- $q_1$  distributes as a  $\chi^2$  under  $s(m_H)+b$  experiments ( $H_1$ )

- The exclusion significance  
 $Z = \sqrt{q_1} = \sqrt{-2 \ln \lambda(1)}$   
can be expressed in terms of an equivalent exclusion CL

$$p_1 = p_{s+b} = 1 - CL$$

- The exclusion sensitivity is the median CL, and using toy MCs one can find the 1 and 2  $\sigma$  bands



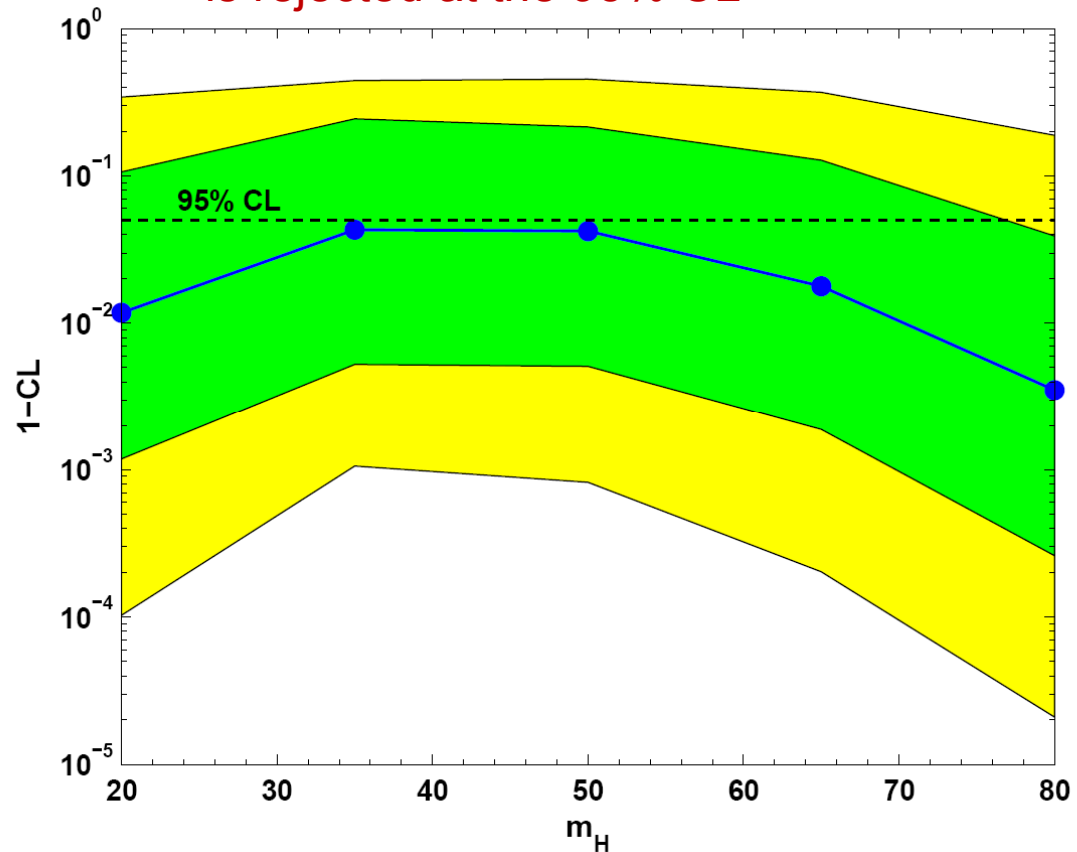
# Exclusion Profile Likelihood Ratio

- A Higgs with a specific mass  $m_H$  is excluded at the 95% CL if the observed p-value of the  $s(m_H)+b$  hypothesis is below 0.05

$$p_1 = p_{s+b} = 1 - CL$$

- In this example a Higgs Boson is expected to be excluded  $p_1 < 0.05$  (CL > 95%) in all the mass range

If  $p_{s+b} < 5\%$ , the  $s(m_H)+b$  hypothesis is rejected at the 95% CL



# Exclusion Bayesian

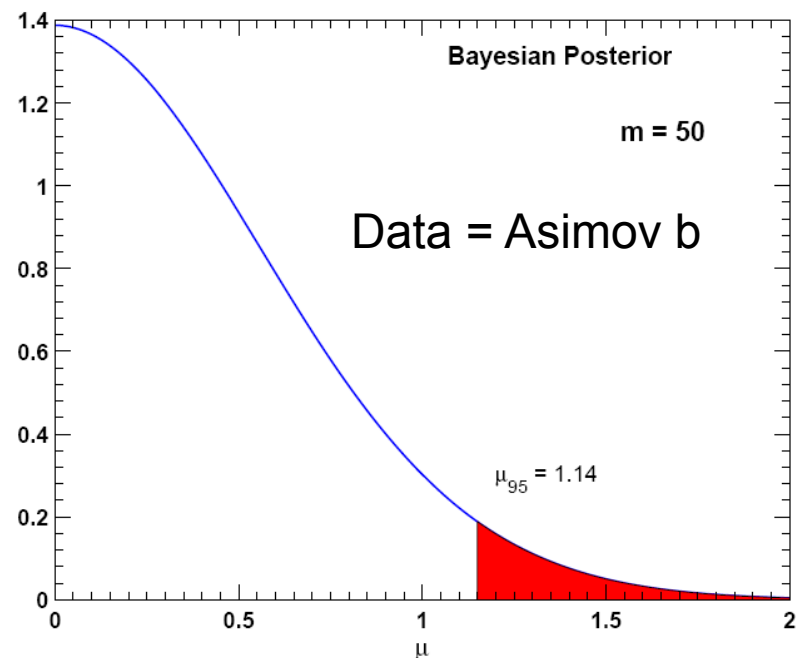
Let  $prob(\mu | n, m)$  be the posterior for  $\mu$

$$prob(\mu | n, m) = \frac{\int L(n, m | \mu \cdot s + b(\theta)) \pi(\mu) \pi(\theta) d\theta}{\iint L(n, m | \mu \cdot s + b(\theta)) \pi(\mu) \pi(\theta) d\theta d\mu}$$

**NOTE:** The pdf of the posterior is based on the **one** observed data event with the likelihood integrated over the nuisance parameters

To set an upper limit on the signal strength  $\mu = \frac{\sigma}{\sigma_{SM}}$  calculate the credibility interval  $[0, \mu_{95}]$

$$0.95 = \int_0^{\mu_{95}} Prob(\mu | n, m) d\mu$$





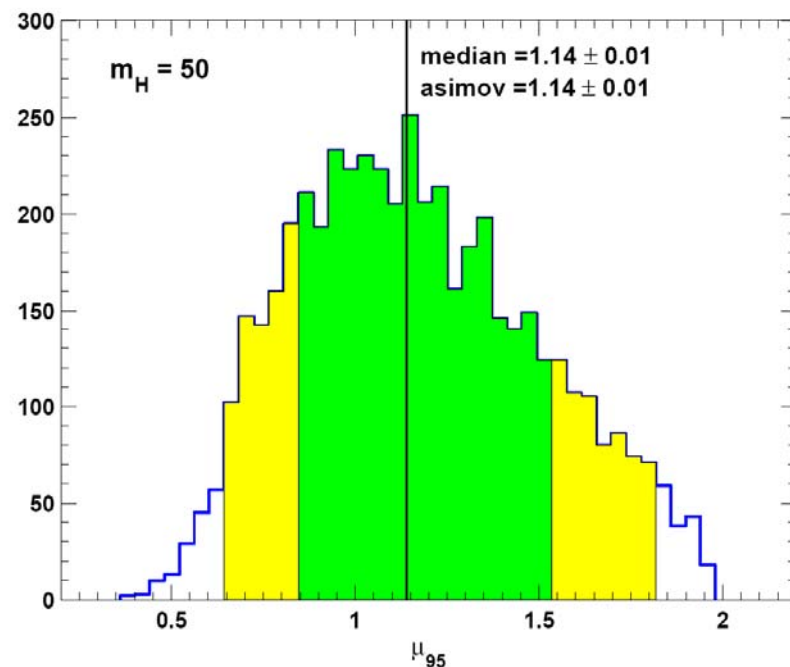
# Exclusion Bayesian

Let  $prob(\mu | n, m)$  be the posterior for  $\mu$

**NOTE:** The toy MC are needed just to find the median sensitivity, but once the data is delivered, it is sufficient to determine the upper limit using the posterior integration

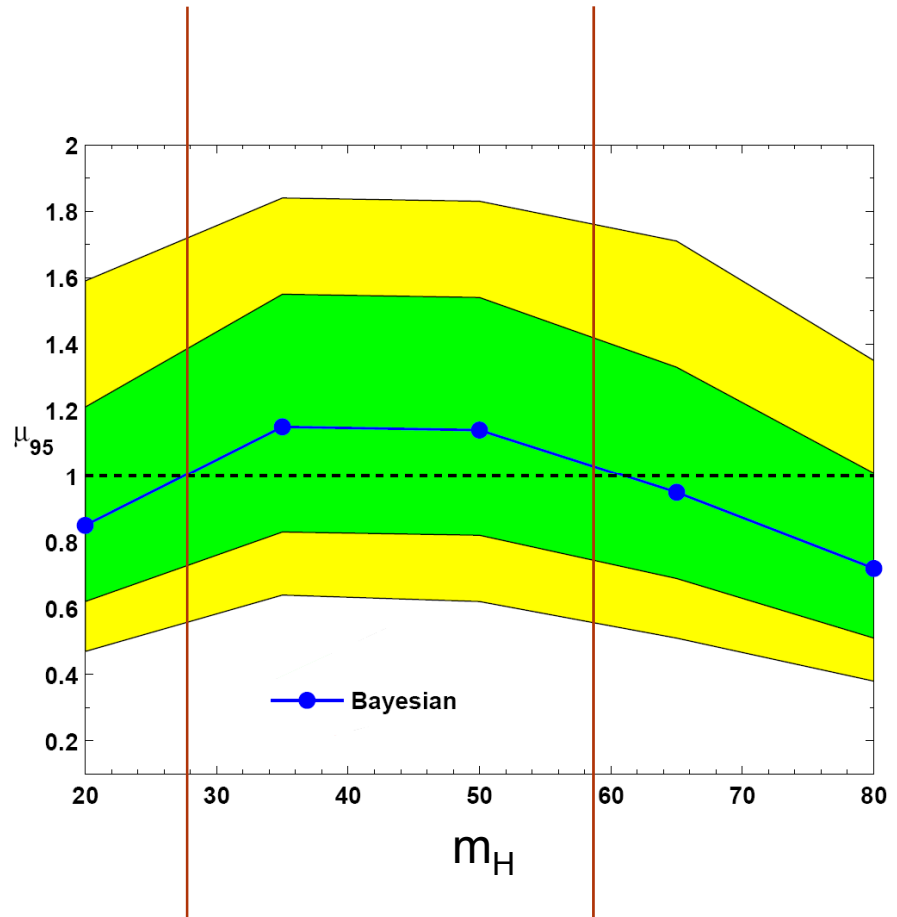
$$0.95 = \int_0^{\mu_{95}} Prob(\mu | n, m) d\mu$$

Data = b-only



# Exclusion Bayesian

- We find that the credibility interval  $[0, \mu_{95}]$  does not contain  $\mu_{95}=1$  (SM) for  $m_H < 28$  or  $m_H > 61$
- This is sometimes **wrongly** expressed as an exclusion at the 95% CL



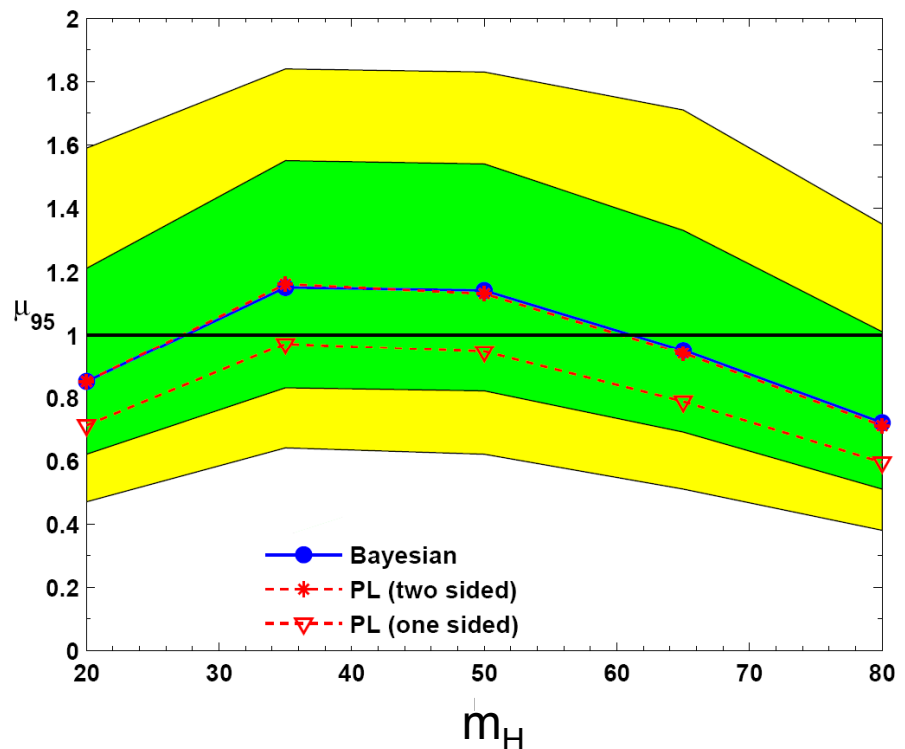
# Exclusion Bayesian vs PL Ratio

- Comparing a credibility Bayesian interval to 95% frequentist CL is like comparing oranges to apples.... yet
- The saddle point approximation

$$\text{prob}(\mu | x) \sim \lambda(\mu) \sim e^{-\frac{\mu^2}{2\sigma^2}}$$

ensures that the Profile Likelihood Ratio and the marginalized Bayes limits are equivalent in this flat priors example

- See appendix for the proof

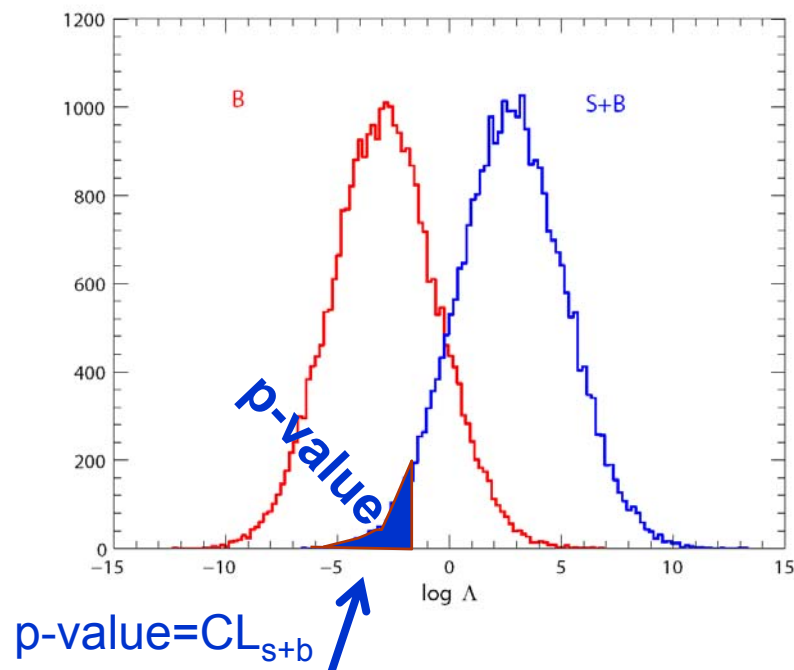


• **NOTE:** One has to be careful about the 1-sided vs 2-sided significance



# The frequentist $CL_{s+b}$ method

- Use the LR as a test statistics  $\Lambda = \frac{L(H_1)}{L(H_0)} = \frac{L(n, m | s + b(\theta))}{L(n, m | b(\theta))}$
- To take systematics into account integrate the nuisance parameters or profile them
- The exclusion is given by the  $s(m_H)+b$  hypothesis p-value  
 $p_{s+b} = CL_{s+b}$
- If  $p_{s+b} < 5\%$ , the  $s(m_H)+b$  hypothesis is rejected at the 95% CL



# The modified frequentist CLs

- $CL_{s+b}$  enables the exclusion of the  $s(m_H)+b$  hypothesis, a downward fluctuation of the background might lead to an exclusion of a signal to which one is not sensitive (with a very low cross section)
- To protect against such fluctuations, the CL was redefined in a non-frequentist way to be

$$CL_s = \frac{CL_{s+b}}{CL_b} = \frac{p_{s+b}}{1 - p_b} \sim \frac{p(n_{obs} \leq s + b)}{p(n_{obs} \leq b)}$$

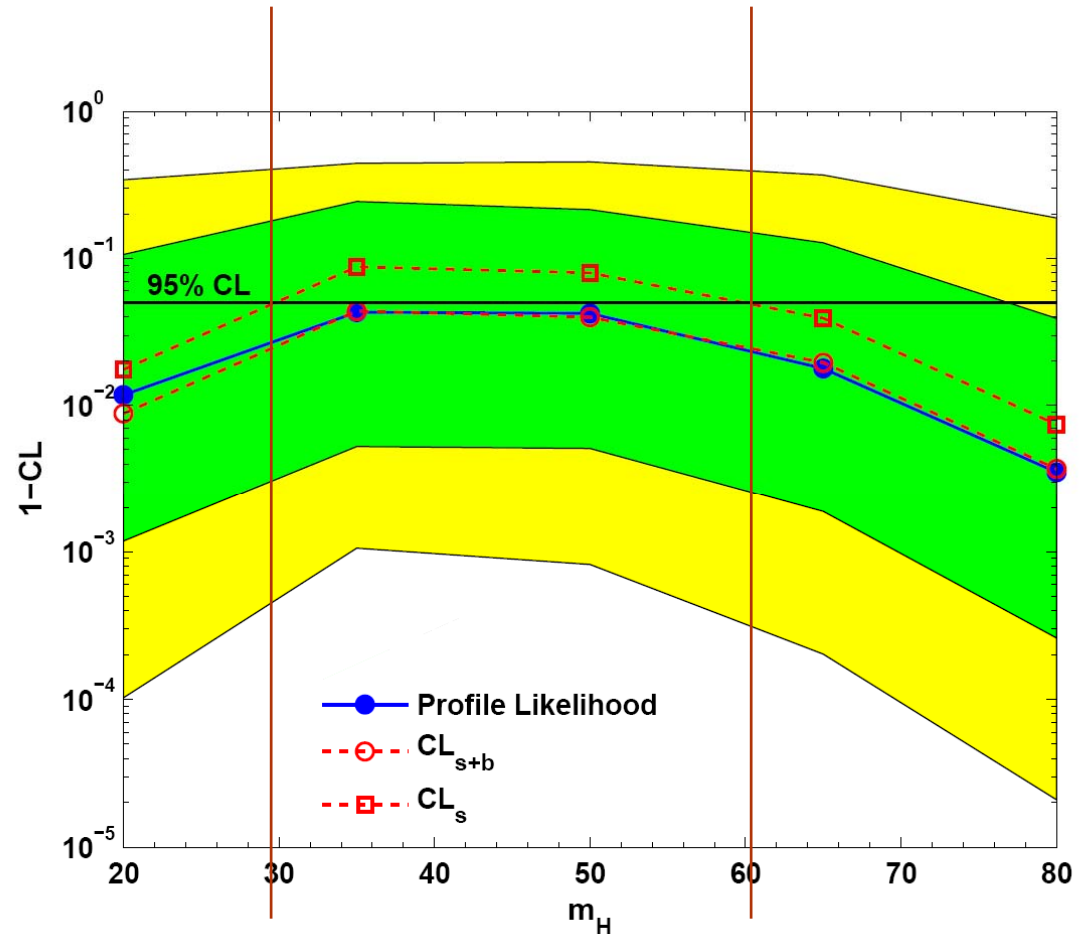
Alex Read  
J.Phys.G28:2693-2704,2002

- Statisticians do not like this p-values ratio, yet, physics-wise it is conservative in a sense of coverage.



# The modified frequentist CLs

- In this example, while using PL or the CLs the Higgs is excluded in all the mass range, the CLs reduces the sensitivity and does not allow to exclude a Higgs with  $30 < m_H < 60$



# Conclusions

- We have explored and compared all the methods to test hypotheses that are currently in use in the High Energy Physics market (PLR,  $CL_{s+b}$ ,  $CL_s$ , Bayesian )
- We have shown a way to appreciate the Bayes factor by comparing it to the PLR
- We have shown that all methods tend to give similar results, (for both exclusion and discovery using flat priors) whether one integrates the nuisance parameters or profile them
- Even though we have used typical case studies, real life might be different and all available methods should be explored



# APPENDIX & BACKUP





# Discovery vs Exclusion

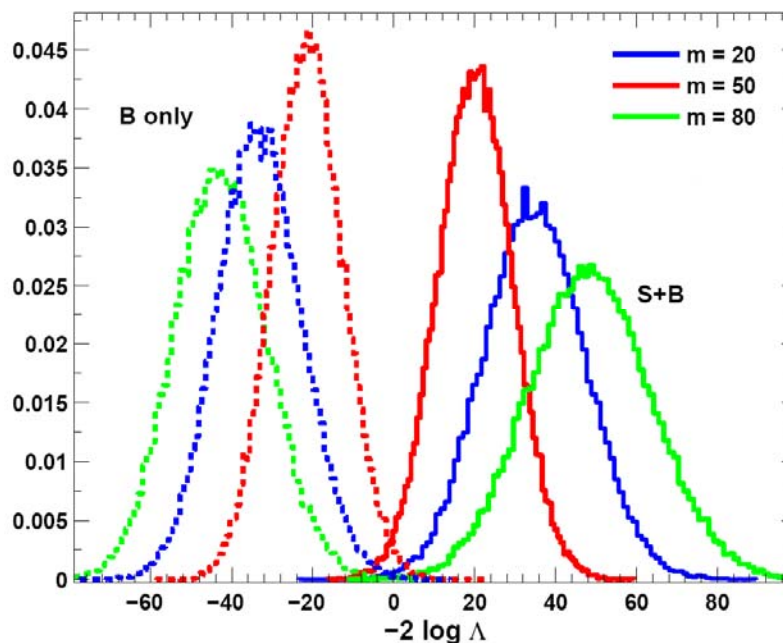
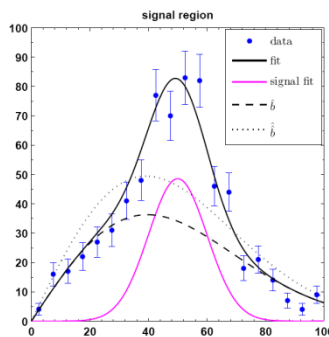
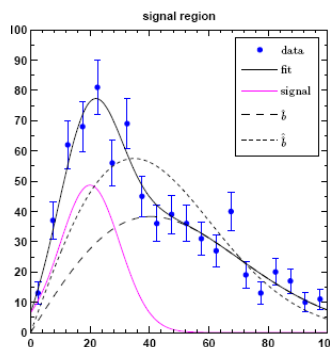
- Higgs statistics is about testing one hypothesis against another hypothesis
  - One hypothesis is the Standard Model with no Higgs Boson ( $H_0$ )
  - Another hypothesis is the SM with a Higgs boson with a specific mass  $m_H$  ( $H_1$ )
- Rejecting the No-Higgs ( $H_0$ ) hypothesis  $\rightarrow$  DISCOVERY
- Rejecting the Higgs hypothesis ( $H_1$ )  $\rightarrow$  EXCLUDING the Higgs
- **The null hypothesis** is the hypothesis we try to reject in favor of **the alternate hypothesis**.
  - DISCOVERY: Null= $H_0$ , Alternate= $H_1$
  - EXCLUSION: Null= $H_1$ , Alternate= $H_0$



# The Profiled CL way


- The signal close to the end of the background mass spectrum is better separated than the signal in the middle

$$\Lambda_{PL} = \frac{L(H_1)}{L(H_0)} = \frac{L\left(n, m \mid s + b(\hat{\theta}_{s+b})\right)}{L\left(n, m \mid b(\hat{\theta}_b)\right)}$$



# PL vs Bayesian

$$\begin{aligned}
 B_{10}(\mu) &= \frac{\int L(\mu \cdot s + b) \pi(\mu) \pi(b) db}{\int L((\mu=0) \cdot s + b) \pi(b) db} \\
 &= \frac{\int e^{\log L(\mu \cdot s + b) \pi(\mu) \pi(b)} db}{\int e^{\log L((\mu=0) \cdot s + b) \pi(b)} db} \\
 &\approx \frac{e^{\log L(\mu \cdot s + \hat{b}(\mu))}}{e^{\log L((\mu=0) \cdot s + \hat{b}(0))}} = \frac{\lambda(\mu)}{\lambda(0)}
 \end{aligned}$$




saddle-point approximation (for flat priors)

[http://en.wikipedia.org/wiki/Method\\_of\\_steepest\\_descent](http://en.wikipedia.org/wiki/Method_of_steepest_descent)



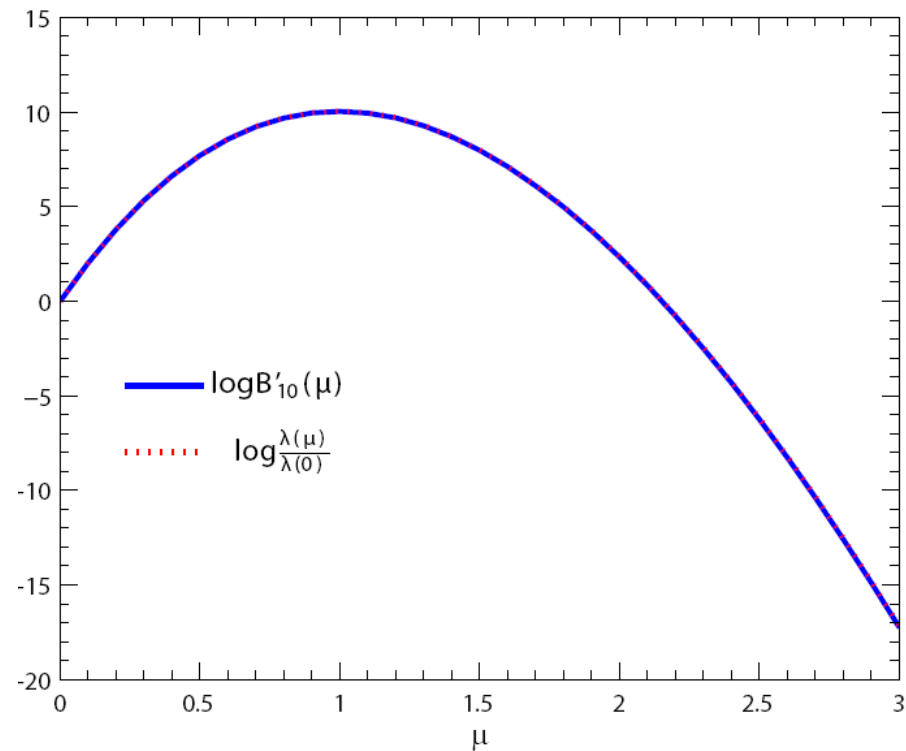
# PL vs Bayesian

$$\begin{aligned}
 B_{10}(\mu) &= \frac{\int L(\mu \cdot s + b) \pi(\mu) \pi(b) db}{\int L((\mu=0) \cdot s + b) \pi(b) db} \\
 &= \frac{\int e^{\log L(\mu \cdot s + b) \pi(\mu) \pi(b)} db}{\int e^{\log L((\mu=0) \cdot s + b) \pi(b)} db} \\
 &\approx \frac{e^{\log L(\mu \cdot s + \hat{b}(\mu))}}{e^{\log L((\mu=0) \cdot s + \hat{b}(0))}} = \frac{\lambda(\mu)}{\lambda(0)}
 \end{aligned}$$



saddle-point approximation

[http://en.wikipedia.org/wiki/Method\\_of\\_steepest\\_descent](http://en.wikipedia.org/wiki/Method_of_steepest_descent)



Asimov Data = S+B



# PL vs Bayesian ?

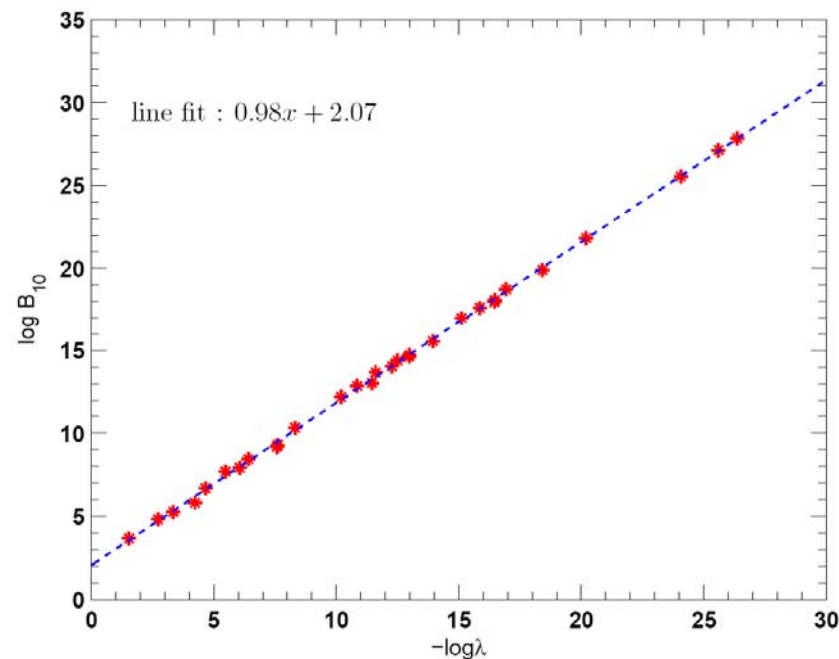
$$B_{10}(\mu) \approx \frac{\lambda(\mu)}{\lambda(0)}$$

$$B_{10} = \int B_{10}(\mu) d\mu$$

$$B_{10} \approx \frac{1}{\lambda(0)} \int \lambda(\mu) d\mu$$

$$\log B_{10} \approx -\log \lambda(0) + \log \int \lambda(\mu) d\mu$$

$$\log B_{10} \approx -\log \lambda(0) + C$$



# PL vs Bayesian ?

$$B'_{10}(\mu) \approx \frac{\lambda(\mu)}{\lambda(0)}$$

$$B_{10} = \int B'_{10}(\mu) d\mu$$

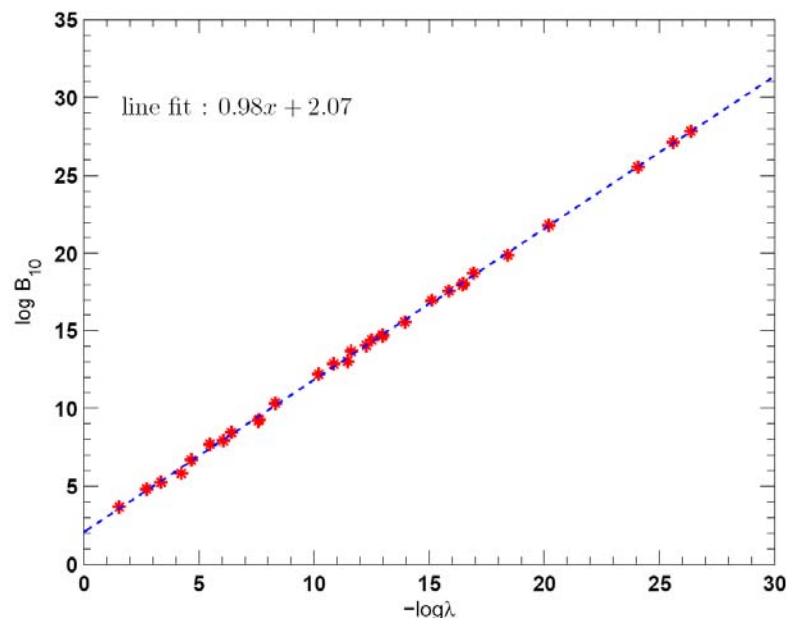
$$B_{10} \approx \frac{1}{\lambda(0)} \int \lambda(\mu) d\mu$$

$$\log B_{10} \approx -\log \lambda(0) + \log \int \lambda(\mu) d\mu$$

$$\log B_{10} \approx -\log \lambda(0) + C$$

$$Z = \sqrt{-2 \log \lambda(0)} \Rightarrow \frac{Z^2}{2} = -\log \lambda(0)$$

$$\log B_{10} \approx \frac{Z^2}{2} + C$$



# Comparing Bayesian to Frequentist PL

$$prob(\mu | \vec{n}, \vec{m}) = \frac{\int L(n, m | \mu \cdot s + b(\theta)) \pi(\mu) \pi(\theta) d\theta}{\iint L(n, m | \mu \cdot s + b(\theta)) \pi(\mu) \pi(\theta) d\theta d\mu} =$$

saddle-point approximation  
(for flat priors)

$$= \frac{e^{\log L(\mu s + \hat{b})}}{e^{\log L(\hat{\mu} s + \hat{b})}} = e^{\log \frac{L(\mu s + \hat{b})}{L(\hat{\mu} s + \hat{b})}} = \lambda(\mu)$$

$$-2 \log \lambda(\mu) \sim \frac{\mu^2}{\sigma^2}$$

$$prob(\mu | x) \sim \lambda(\mu) \sim e^{-\frac{\mu^2}{2\sigma^2}}$$

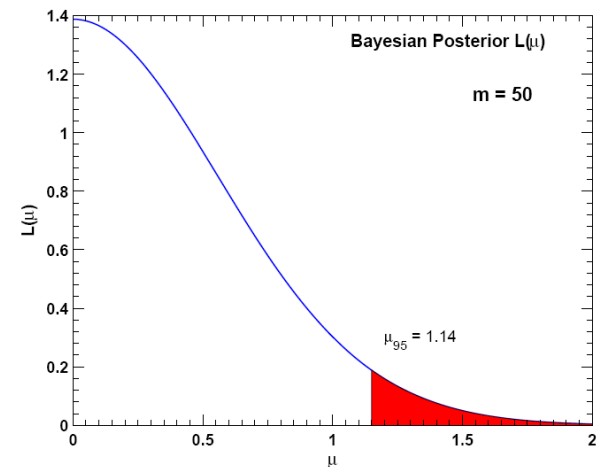
**Note:** 95% Credibility Interval  $\sim$  Gaussian  $1.96\sigma$

The 95% CL Profile Likelihood limit is based on  $1.64\sigma$  (2-sided 10% ~ 1 Sided 5%)

When comparing the expected limits

(though it is like comparing oranges to apples)

one has to be careful about the definitions



# Wilks Theorem

S.S. Wilks, *The large-sample distribution of the likelihood ratio for testing composite hypotheses*, Ann. Math. Statist. **9** (1938) 60-2.

$$\lambda(\mu = 0) = \frac{L(0 \cdot s(m_H) + \hat{\hat{b}} \mid x)}{L(\hat{\mu} \cdot s(m_H) + \hat{\hat{b}} \mid x)}$$

- Under a set of regularity conditions and for a sufficiently large data sample, Wilks' theorem says that for a hypothesized value of  $\mu=0$ , the pdf of the statistic  $q_0 = -2\ln\lambda(\mu=0)$  approaches the chi-square pdf for one degree of freedom





# PL Discovery - Illustrated

*The profile LR of bg - only experiments ( $\mu = 0$ )  
under the hypothesis of BG only ( $H_0$ )*

$$f(q_0 | \mu = 0)$$

*The profile LR of  $S + B$  experiments ( $\mu = 1$ )  
under the hypothesis of BG only ( $H_0$ )*

$$f(q_0 | \mu = 1)$$

*The observed profile LR*

$$q_{0,obs} = -2 \log \frac{L(0 \cdot s + b | data)}{L(\hat{\mu} \cdot s + b | data)}$$

$$p_0 = \int_{q_{0,obs}}^{\infty} f(q_0 | 0) dq_0$$

$p_0$  is the level of compatibility between the data and the no-Higgs hypothesis  
If  $p_0$  is smaller than  $\sim 2.8 \cdot 10^{-7}$  we claim a  $5\sigma$  discovery

Elia Grosse, Higgs Statistics, Gracow 2009

# PL Exclusion - Illustrated

The profile LR of  $s + b$  experiments ( $\mu = 1$ )

under the hypothesis of  $s + b$  ( $H_1$ )

$$f(q_1 | \mu = 1)$$

The profile LR of  $b$ -only experiments ( $\mu = 0$ )

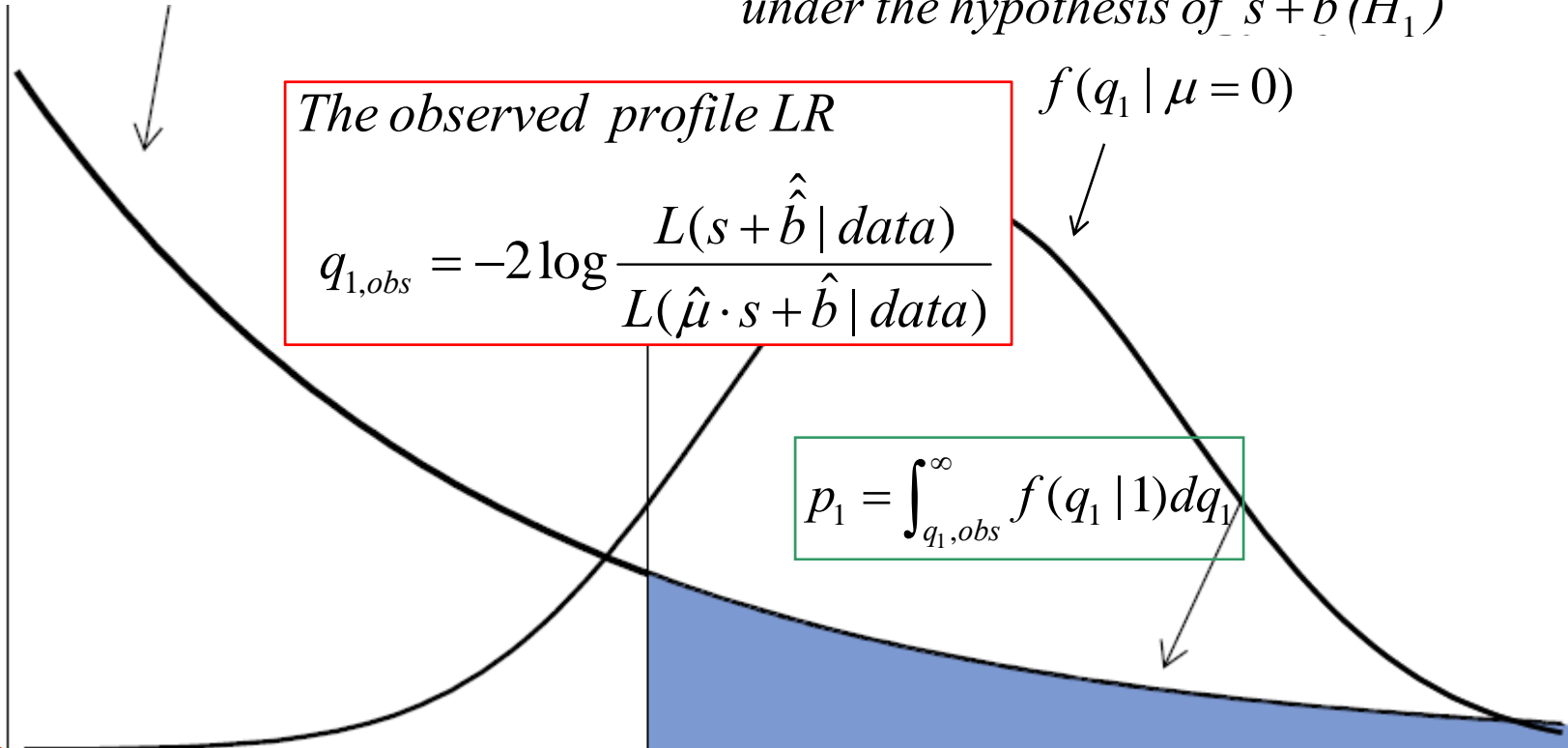
under the hypothesis of  $s + b$  ( $H_1$ )

$$f(q_1 | \mu = 0)$$

The observed profile LR

$$q_{1,obs} = -2 \log \frac{L(s + \hat{b} | data)}{L(\hat{\mu} \cdot s + \hat{b} | data)}$$

$$p_1 = \int_{q_{1,obs}}^{\infty} f(q_1 | 1) dq_1$$



Filippos Gross, Higgs Statistics, Cracow 2009

$p_1$  is the level of compatibility between the data and the Higgs hypothesis

If  $p_1$  is smaller than 0.05 we claim an exclusion at the 95% CL